

# Standard Practice for Applying Statistical Quality Assurance Techniques to Evaluate Analytical Measurement System Performance<sup>1</sup>

This standard is issued under the fixed designation D 6299; the number immediately following the designation indicates the year of original adoption or, in the case of revision, the year of last revision. A number in parentheses indicates the year of last reapproval. A superscript epsilon ( $\epsilon$ ) indicates an editorial change since the last revision or reapproval.

#### 1. Scope

1.1 This practice provides information for the design and operation of a program to monitor and control ongoing stability and precision and bias performance of selected analytical measurement systems using a collection of generally accepted statistical quality control (SQC) procedures and tools.

NOTE 1—A complete list of criteria for selecting measurement systems to which this practice should be applied and for determining the frequency at which it should be applied is beyond the scope of this practice. However, some factors to be considered include (I) frequency of use of the analytical measurement system, (2) criticality of the parameter being measured, (3) system stability and precision performance based on historical data, (4) business economics, and (5) regulatory, contractual, or test method requirements.

1.2 This practice is applicable to stable analytical measurement systems that produce results on a continuous numerical scale.

1.3 This practice is applicable to laboratory test methods.

1.4 This practice is applicable to validated process stream analyzers.

Note 2—For validation of univariate process stream analyzers, see also Practice D 3764.

1.5 This practice assumes that the normal (Gaussian) model is adequate for the description and prediction of measurement system behavior when it is in a state of statistical control.

NOTE 3—For non-Gaussian processes, transformations of test results may permit proper application of these tools. Consult a statistician for further guidance and information.

1.6 This practice does not address statistical techniques for comparing two or more analytical measurement systems applying different analytical techniques or equipment components that purport to measure the same property(s).

# 2. Referenced Documents

2.1 ASTM Standards:

D 3764 Practice for Validation of Process Stream Analyzer Systems<sup>2</sup>

- D 5191 Test Method for Vapor Pressure of Petroleum Products (Mini Method)<sup>2</sup>
- $E\ 177\ Practice \ for \ Use \ of \ the \ Terms \ Precision \ and \ Bias \ in \ ASTM \ Test \ Methods^3$
- E 178 Practice for Dealing with Outlying Observations<sup>3</sup>
- E 456 Terminology Relating to Quality and Statistics<sup>3</sup>
- E 691 Practice for Conducting an Interlaboratory Study to Determine the Precision of a Test Method<sup>3</sup>

#### 3. Terminology

3.1 Definitions:

3.1.1 accepted reference value, n—a value that serves as an agreed-upon reference for comparison and that is derived as (1) a theoretical or established value, based on scientific principles, (2) an assigned value, based on experimental work of some national or international organization, such as the U.S. National Institute of Standards and Technology (NIST), or (3) a consensus value, based on collaborative experimental work under the auspices of a scientific or engineering group.

#### (E 456/E 177)

3.1.2 *accuracy*, *n*—the closeness of agreement between an observed value and an accepted reference value. (E 456/ E 177)

3.1.3 assignable cause, n—a factor that contributes to variation and that is feasible to detect and identify. (E 456)

3.1.4 *bias*, n—a systematic error that contributes to the difference between a population mean of the measurements or test results and an accepted reference or true value. (E 456/ E 177)

3.1.5 *control limits*, *n*—limits on a control chart that are used as criteria for signaling the need for action or for judging whether a set of data does or does not indicate a state of statistical control. (E 456)

3.1.6 *lot*, n—a definite quantity of a product or material accumulated under conditions that are considered uniform for sampling purposes. (E 456)

3.1.7 precision, n-the closeness of agreement between test

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<sup>&</sup>lt;sup>2</sup> Annual Book of ASTM Standards, Vol 05.02.

<sup>&</sup>lt;sup>3</sup> Annual Book of ASTM Standards, Vol 14.02.

results obtained under prescribed conditions.

(E 456)

3.1.8 *repeatability conditions, n*—conditions where mutually independent test results are obtained with the same test method in the same laboratory by the same operator with the same equipment within short intervals of time, using test specimens taken at random from a single sample of material. (E 456, E 177)

3.1.9 *reproducibility conditions*, *n*—conditions under which test results are obtained in different laboratories with the same test method, using test specimens taken at random from the same sample of material. (E 456, E 177)

3.2 Definitions of Terms Specific to This Standard:

3.2.1 analytical measurement system, n—a collection of one or more components or subsystems, such as samplers, test equipment, instrumentation, display devices, data handlers, printouts or output transmitters, that is used to determine a quantitative value of a specific property for an unknown sample in accordance with a test method.

3.2.1.1 *Discussion*—An analytical measurement system may comprise multiple instruments being used for the same test method.

3.2.2 *blind submission*, *n*—submission of a check standard or quality control (QC) sample for analysis without revealing the expected value to the person performing the analysis.

3.2.3 *check standard*, *n*—*in QC testing*, a material having an accepted reference value used to determine the accuracy of a measurement system.

3.2.3.1 *Discussion*—A check standard is preferably a material that is either a certified reference material with traceability to a nationally recognized body or a material that has an accepted reference value established through interlaboratory testing. For some measurement systems, a pure, single component material having known value or a simple gravimetric or volumetric mixture of pure components having calculable value may serve as a check standard. Users should be aware that for measurement systems that show matrix dependencies, accuracy determined from pure compounds or simple mixtures may not be representative of that achieved on actual samples.

3.2.4 *common (chance, random) cause, n*—for quality assurance programs, one of generally numerous factors, individually of relatively small importance, that contributes to variation, and that is not feasible to detect and identify.

3.2.5 *double blind submission*, *n*—submission of a check standard or QC sample for analysis without revealing the check standard or QC sample status and expected value to the person performing the analysis.

3.2.6 *expected value*, *n*—for a QC sample analyzed using an in-statistical control measurement system, the estimate of the theoretical limiting value to which the average of results tends when the number of results approaches infinity.

3.2.7 *in-statistical-control*, *adj*—a process, analytical measurement system, or function that exhibits variations that can only be attributable to common cause.

3.2.8 *proficiency testing*, *n*—determination of a laboratory's testing capability by participation in an interlaboratory cross-check program.

3.2.8.1 *Discussion*—ASTM Committee D02 conducts proficiency testing among hundreds of laboratories, using a wide variety of petroleum products and lubricants.

3.2.9 quality control (QC) sample, n—for use in quality assurance programs to determine and monitor the precision and stability of a measurement system, a stable and homogeneous material having physical or chemical properties, or both, similar to those of typical samples tested by the analytical measurement system. The material is properly stored to ensure sample integrity, and is available in sufficient quantity for repeated, long term testing.

3.2.10 site precision (R'), *n*—the value below which the absolute difference between two individual test results obtained under site precision conditions may be expected to occur with a probability of approximately 0.95 (95 %). It is defined as 2.77 times the standard deviation of results obtained under site precision conditions.

3.2.11 site precision conditions, n—conditions under which test results are obtained by one or more operators in a single site location practicing the same test method on a single measurement system which may comprise multiple instruments, using test specimens taken at random from the same sample of material, over an extended period of time spanning at least a 15 day interval.

3.2.11.1 *Discussion*—Site precision conditions should include all sources of variation that are typically encountered during normal, long term operation of the measurement system. Thus, all operators who are involved in the routine use of the measurement system should contribute results to the site precision determination. If multiple results are obtained within a 24–h period, then it is recommended that the number of results used in site precision calculations be increased to capture the longer term variation in the system.

3.2.12 site precision standard deviation, n—the standard deviation of results obtained under site precision conditions.

3.2.13 validation audit sample, n—a QC sample or check standard used to verify precision and bias estimated from routine quality assurance testing.

3.3 Symbols:

3.3.1 ARV—accepted reference value.

3.3.2 EWMA—exponentially weighted moving average.

3.3.3 *I*—individual observation (as in *I*-chart).

3.3.4 *MR*—moving range.

3.3.5  $\overline{MR}$  —average of moving range.

3.3.6 QC—quality control.

3.3.7 R'—site precision.

3.3.8  $\sigma_{R'}$ —site precision standard deviation.

3.3.9 VA—validation audit.

3.3.10  $\chi^2$ —chi squared.

3.3.11  $\lambda$ —lambda.

#### 4. Summary of Practice

4.1 QC samples and check standards are regularly analyzed by the measurement system. Control charts and other statistical techniques are presented to screen, plot, and interpret test results in accordance with industry-accepted practices to ascertain the in-statistical-control status of the measurement system.

4.2 Statistical estimates of the measurement system precision and bias are calculated and periodically updated using accrued data. 4.3 In addition, as part of a separate validation audit procedure, QC samples and check standards may be submitted blind or double-blind and randomly to the measurement system for routine testing to verify that the calculated precision and bias are representative of routine measurement system performance when there is no prior knowledge of the expected value or sample status.

#### 5. Significance and Use

5.1 This practice can be used to continuously demonstrate the proficiency of analytical measurement systems that are used for establishing and ensuring the quality of petroleum and petroleum products.

5.2 Data accrued, using the techniques included in this practice, provide the ability to monitor analytical measurement system precision and bias.

5.3 These data are useful for updating test methods as well as for indicating areas of potential measurement system improvement.

#### 6. Reference Materials

6.1 QC samples are used to establish and monitor the precision of the analytical measurement system.

6.1.1 Select a stable and homogeneous material having physical or chemical properties, or both, similar to those of typical samples tested by the analytical measurement system.

NOTE 4—When the QC sample is to be utilized for monitoring a process stream analyzer performance, it is often helpful to supplement the process analyzer system with a subsystem to automate the extraction, mixing, storage, and delivery functions associated with the QC sample.

6.1.2 Estimate the quantity of the material needed for each specific lot of QC sample to (1) accommodate the number of analytical measurement systems for which it is to be used (laboratory test apparatuses as well as process stream analyzer systems) and (2) provide determination of QC statistics for a useful and desirable period of time.

6.1.3 Collect the material into a single container and isolate it.

6.1.4 Thoroughly mix the material to ensure homogeneity.

6.1.5 Conduct any testing necessary to ensure that the QC sample meets the characteristics for its intended use.

6.1.6 Package or store OC samples, or both, as appropriate for the specific analytical measurement system to ensure that all analyses of samples from a given lot are performed on essentially identical material. If necessary, split the bulk material collected in 6.1.3 into separate and smaller containers to help ensure integrity over time. (Warning-Treat the material appropriately to ensure its stability, integrity, and homogeneity over the time period for which it is to stored and used. For samples that are volatile, such as gasoline, storage in one large container that is repeatedly opened and closed can result in loss of light ends. This problem can be avoided by chilling and splitting the bulk sample into smaller containers, each with a quantity sufficient to conduct the analysis. Similarly, samples prone to oxidation can benefit from splitting the bulk sample into smaller containers that can be blanketed with an inert gas prior to being sealed and leaving them sealed until the sample is needed.)

analytical measurement system.

6.2.1 A check standard may be a commercial standard reference material when such material is available in appropriate quantity, quality and composition.

NOTE 5—Commercial reference material of appropriate composition may not be available for all measurement systems.

6.2.2 Alternatively, a check standard may be prepared from a material that is analyzed under reproducibility conditions by multiple measurement systems. The accepted reference value (ARV) for this check standard shall be the average after statistical examination and outlier treatment has been applied.<sup>4</sup>

6.2.2.1 Exchange samples circulated as part of an interlaboratory exchange program, or round robin, may be used as check standards. For an exchange sample to be usable as a check standard, the standard deviation of the interlaboratory exchange program shall not be statistically greater than the reproducibility standard deviation for the test method. An *F*-test should be applied to test acceptability.

NOTE 6—The uncertainty in the ARV is inversely proportional to the square root of the number of values in the average. This practice recommends that a minimum of 16 non-outlier results be used in calculating the ARV to reduce the uncertainty of the ARV by a factor of 4 relative to the measurement system single value precision. The bias tests described in this practice assume that the uncertainty in the ARV is negligible relative to the measurement system precision. If less than 16 values are used in calculating the average, this assumption may not be valid.

NOTE 7—Examples of exchanges that may be acceptable are ASTM D02.CS92 ILCP program; ASTM D02.01 N.E.G.; ASTM D02.01.A Regional Exchanges; International Quality Assurance Exchange Program, administered by Alberta Research Council.

6.2.3 For some measurement systems, single, pure component materials with known value, or simple gravimetric or volumetric mixtures of pure components having calculable value may serve as a check standard. For example, pure solvents, such as 2,2-dimethylbutane, are used as check standards for the measurement of Reid vapor pressure by Test Method D 5191. Users should be aware that for measurement systems that show matrix dependencies, accuracy determined from pure compounds or simple mixtures may not be representative of that achieved on actual samples.

6.3 Validation audit (VA) samples are QC samples and check standards, which may, at the option of the users, be submitted to the measurement system in a blind, or double blind, and random fashion to verify precision and bias estimated from routine quality assurance testing.

#### 7. Quality Assurance (QA) Program for Individual Measurement Systems

7.1 *Overview*—A QA program  $(1)^5$  can consist of five primary activities: (1) monitoring stability and precision through QC sample testing, (2) monitoring accuracy, (3)

6.2 Check standards are used to estimate the accuracy of the

<sup>&</sup>lt;sup>4</sup> Refer to Research Report RR:D02–1007 and Practices E 178 and E 691 in *ASTM Standards on Precision and Bias for Various Applications*, ASTM International, for guidance in statistical and outlier treatment of data. Request PCN:03-512088-34.

<sup>&</sup>lt;sup>5</sup> The boldface numbers in parentheses refer to the list of references at the end of this standard.

periodic evaluation of system performance in terms of precision or bias, or both, (4) proficiency testing through participation in interlaboratory exchange programs where such programs are available, and (5) a periodic and independent system validation using VA samples may be conducted to provide additional assurance of the system precision and bias metrics established from the primary testing activities. At minimum, the QA program must include at least item one.

NOTE 8—For some measurement systems, suitable check standard materials may not exist, and there may be no reasonably available exchange programs to generate them. For such systems, there is no means of verifying the accuracy of the system, and the QA program will only involve monitoring stability and precision through QC sample testing.

7.2 Monitoring System Stability and Precision Through QC Sample Testing—QC test specimen samples from a specific lot are introduced and tested in the analytical measurement system on a regular basis to establish system performance history in terms of both stability and precision.

7.3 Monitoring Accuracy:

7.3.1 Check standards can be tested in the analytical measurement system on a regular basis to establish system performance history in terms of accuracy.

7.3.2 For measurement systems where calibration is established by using multiple standards of known values, such as materials certified by or traceable to the national certification bodies such as NIST, JIS, BSI, and so forth, and where the total number of standards used exceed the number of parameters estimated by the calibration equation, an alternative approach (instead of check standard testing) to infer system accuracy is to compare the statistics associated with the calibration equation to previously established measurement system precision and to standard errors of the calibration standards used. Coverage of this type of statistical techniques for accuracy inference is beyond the scope of this practice. Users are advised to enlist the services of a statistician when using this approach to infer system accuracy instead of check standard testing.

7.4 Test Program Conditions/Frequency:

7.4.1 Conduct both QC sample and check standard testing under site precision conditions.

NOTE 9—It is inappropriate to use test data collected under repeatability conditions to estimate the long term precision achievable by the site because the majority of the long term measurement system variance is due to common cause variations associated with the combination of time, operator, reagents, instrumentation calibration factors, and so forth, which would not be observable in data obtained under repeatability conditions.

7.4.2 Test the QC and check standard samples on a regular schedule, as appropriate. Principal factors to be considered for determining the frequency of testing are (1) frequency of use of the analytical measurement system, (2) criticality of the parameter being measured, (3) established system stability and precision performance based on historical data, (4) business economics, and (5) regulatory, contractual, or test method requirements.

NOTE 10—At the discretion of the laboratory, check standards may be used as QC samples. In this case, the results for the check standards may be used to monitor both stability (see 7.2) and accuracy (see 7.3) simultaneously. If check standards are expensive, or not available in

sufficient quantity, then separate QC samples are employed. In this case, the accuracy (see 7.3) is monitored less frequently, and the QC sample testing (see 7.2) is used to demonstrate the stability of the measurement system between accuracy tests.

7.4.3 It is recommended that a QC sample be analyzed at the beginning of any set of measurements and immediately after a change is made to the measurement system.

7.4.4 Establish a protocol for testing so that all persons who routinely operate the system participate in generating QC test data.

7.4.5 Handle and test the QC and check standard samples in the same manner and under the same conditions as samples or materials routinely analyzed by the analytical measurement system.

7.4.6 When practical, randomize the time of check standard and additional QC sample testing over the normal hours of measurement system operation, unless otherwise prescribed in the specific test method.

NOTE 11—Avoid special treatment of QC samples designed to get a better result. Special treatment seriously undermines the integrity of precision estimates.

# 7.5 Evaluation of System Performance in Terms of Precision and Bias:

7.5.1 Pretreat and screen results accumulated from QC and check standard testing. Apply statistical techniques to the pretreated data to identify erroneous data. Plot appropriately pretreated data on control charts.

7.5.2 Periodically analyze results from control charts, excluding those data points with assignable causes, to quantify the bias and precision estimates for the measurement system.

7.6 Proficiency Testing:

7.6.1 Participation in regularly conducted interlaboratory exchanges where typical production samples are tested by multiple measurement systems, using a specified (ASTM) test protocol, provide a cost-effective means of assessing measurement system accuracy relative to average industry performance. Such proficiency testing can be used instead of check standard testing for systems where the timeliness of the accuracy check is not critical. Proficiency testing may be used as a supplement to accuracy monitoring by way of check standard testing.

7.6.2 Participants plot their signed deviations from the consensus values (exchange averages) on control charts in the same fashion described below for check standards, to ascertain if their measurement processes are non-biased relative to industry average.

7.7 Independent System Validation—Periodically, at the discretion of users, VA samples may be submitted blind or double blind for analysis. Precision and bias estimates calculated using VA samples test data can be used as an independent validation of the routine QA program performance statistics.

NOTE 12—For measurement systems susceptible to human influence, the precision and bias estimates calculated from data where the analyst is aware of the sample status (QC or check standard) or expected values, or both, may underestimate the precision and bias achievable under routine operation. At the discretion of the users, and depending on the criticality of these measurement systems, the QA program may include periodic blind or double-blind testing of VA samples.

7.7.1 The specific design and approach to the VA testing

program will depend on features specific to the measurement system and organizational requirements, and is beyond the intended scope of this practice. Some possible approaches are noted as follows.

7.7.1.1 If all QC samples or check standards, or both, are submitted blind or double blind and the results are promptly evaluated, then additional VA sample testing may not be necessary.

7.7.1.2 QC samples or check standards, or both, may be submitted as unknown samples at a specific frequency. Such submissions should not be so regular as to compromise their blind status.

7.7.1.3 Retains of previously analyzed samples may be resubmitted as unknown samples under site precision conditions. Generally, data from this approach can only yield precision estimates as retain samples do not have ARVs. Typically, the differences between the replicate analyses are plotted on control charts to estimate the precision of the measurement system. If precision is level dependent, the differences are scaled by the standard deviation of the measurement system precision at the level of the average of the two results.

# 8. Procedure for Pretreatment, Assessment, and Interpretation of Test Results

8.1 *Overview*—Results accumulated from QC, check standard, and VA sample testing are pretreated and screened. Statistical techniques are applied to the pretreated data to achieve the following objectives:

8.1.1 Identify erroneous data,

8.1.2 Assess initial results,

8.1.3 Deploy, interpret and maintain of control charts, and

8.1.4 Quantify long term measurement precision and bias.

NOTE 13—Refer to the annex for examples of the application of the techniques that are discussed below and described in Section 9.

8.2 *Pretreatment of Test Results*—Assessment, control charting, and evaluation are applied only to appropriately pretreated test results. The purpose of pretreatment is to standardize the control chart scales so as to allow for data from multiple check standards to be compared on the same chart.

8.2.1 For QC sample test results, no data pretreatment is typically used since results for different QC samples are generally not plotted on the same chart.

8.2.2 For check standard sample test results, two cases apply, depending on the measurement system precision:

8.2.2.1 *Case 1*—If either (1) all of the check standard test results are from one or more lots of check standard material having the same ARV(s), or (2) the precision of the measurement system is constant across levels, then pretreatment consists of calculating the difference between the test result and the ARV:

$$Pretreated result = test result - ARV(for the sample)$$
(1)

8.2.2.2 *Case* 2—Test results are for multiple lots of check standards with different ARVs, and the precision of the measurement system is known to vary with level,

$$Pretreated result = \frac{[test result - ARV(for the sample)]}{standard deviation at the ARV level}$$
(2)

where the standard deviation at the ARV level is the published reproducibility standard deviation. In the event that no published reproducibility exists and the ARV was established through round robin testing, standard deviations determined from round robin testing may be used.

8.2.2.3 If there is no published reproducibility standard deviation and the ARV was not arrived at by round robin testing, a standard deviation should be determined by users in a technically acceptable manner.

NOTE 14—It is recommended that the method used to determine the standard deviation be developed under the guidance of a statistician.

Note 15—To calculate the reproducibility standard deviations from published reproducibilities, divide the accepted reproducibility value at each level by 2.77.

8.2.3 Pretreatment of results for VA samples is done in the same manner as described in 8.2.1 and 8.2.2.

8.3 Assessment of Initial Results—Assessment techniques are applied to test results collected during the startup phase of or after significant modifications to a measurement system. Perform the following assessment after at least 15 pretreated results have become available. The purpose of this assessment is to ensure that these results are suitable for deployment of control charts (described in A1.4).

NOTE 16—These techniques can also be applied as diagnostic tools to investigate out-of-control situations.

8.3.1 *Screen for Suspicious Results*—Pretreated results should first be visually screened for values that are inconsistent with the remainder of the data set, such as those that could have been caused by transcription errors. Those flagged as suspicious should be investigated. Discarding data at this stage must be supported by evidence gathered from the investigation. If, after discarding suspicious pretreated results there are less than 15 values remaining, collect additional data and start over.

8.3.2 *Screen for Unusual Patterns*—The next step is to examine the pretreated results for non-random patterns such as continuous trending in either direction, unusual clustering, and cycles. One way to do this is to plot the results on a run chart (see A1.3) and examine the plot. If any non-random pattern is detected, investigate for and eliminate the root cause(s). Discard the data set and start the procedure again.

8.3.3 *Test "Normality" Assumption*—For measurement systems with no prior performance history, or as a diagnostic tool, it is useful to test that the results from the measurement are adequately described by a normal distribution. One way to do this is to use a normal probability plot and the Anderson-Darling Statistic (see A1.4). If the results show obvious deviation from normality, then the statistical control charting techniques described are not directly applicable to the measurement system.

NOTE 17—Transformations may lead to normally distributed data, but these techniques are outside the scope of this practice.

8.4 Control Charts (1, 2)—Individual (*I*) and moving range of two (*MR*) control charts are the recommended tools for (*a*) routine recording of QC sample and check standard test results, and (*b*) immediate assessment of the "in statistical control" (3) status of the system that generated the data. Optionally, the exponentially weighted moving average (*EWMA*) (4, 5) may be overlaid on the *I* chart to enhance detection power for small level shifts. NOTE 18—The control charts and statistical techniques described in this practice are chosen for their simplicity and ease of use. It is not the intent of this practice to preclude use of other statistically equivalent or more advanced techniques, or both.

8.4.1 *Construction of Control Charts*—If no obvious unusual patterns are detected from the run charts, and no obvious deviation from normality is detected, proceed with construction of the control charts

8.4.1.1 *MR Chart*—Construct an *MR* plot and examine it for unusual patterns. If no unusual patterns are found in the *MR* plot, calculate and overlay the control limits on the *MR* plot to complete the *MR* chart.

8.4.1.2 *I Chart*—Calculate control limits and overlay them on the "run chart" to produce the *I* chart.

8.4.1.3 *EWMA Overlay*—Optionally, calculate the *EWMA* values and plot them on the *I* chart. Calculate the *EWMA* control limits and overlay them on the *I* chart.

8.4.2 *Control Chart Deployment*—Put these control charts into operation by regularly plotting the pretreated test results on the charts and immediately interpreting the charts.

8.5 Control Chart Interpretation:

8.5.1 Apply control chart rules (see A1.5) to determine if the data supports the hypothesis that the measurement system is under the influence of common causes variation only (in statistical control).

8.5.2 *Investigate Out-of-Control Points in Detail*—Exclude from further data analysis those associated with assignable causes, provided the assignable causes are deemed not to be part of the normal process.

Note 19-All data, regardless of in-control or out-of-control status, needs to be recorded.

8.6 Scenario 1 for Periodic Updating of Control Charts Parameters:

8.6.1 Scenario 1 covers (a) control charts for a QC material where there had been no change in the system, but more data of the same level has been accrued; or (b) control charts for check standard pretreated results.

8.6.2 When a minimum of 15 new in-control data points becomes available, the precision estimate used to calculate the control limits can be updated to incorporate the information from this new data. Update calculations that involve pooling of old and new data sets shall be preceded by an F-test (see A1.8) of sample variances for the new data set versus the existing in-control data set.

8.6.3 If the outcome of the *F*-test is not significant, then the precision estimate is updated by statistically pooling both sample variances. A significant *F*-test should trigger an investigation for assignable causes.

# 8.7 Scenario 2 for Periodic Updating of Control Charts Parameters:

8.7.1 Scenario 2 covers control chart for QC materials where an assignable cause change in the system had occurred due to a change in the level for the QC material. Minor or major differences may exist between QC material batches. Since control limit calculations for the *I* chart require a center value established by the measurement system, a special transition procedure is required to ensure that the center value for a new batch of QC material is established using results

produced by a measurement system that is in statistical control. This practice presents two procedures to be selected at the users' discretion.

8.7.2 Procedure 1–Concurrent Testing:

8.7.2.1 Collect and prepare a new batch of QC material when the current QC material supply remaining can support no more than 20 analyses.

8.7.2.2 Concurrently test and record data for the new material each time a current QC sample is tested. The result for the new material is deemed valid if the measurement process in-control status is validated by the current QC material and control chart.

8.7.2.3 Optionally, to provide an early indication of the status of the new batch of QC material, immediately start a run chart and an *MR* plot for the new material. After five valid results become available for the new material, convert the run chart into an *I* chart with trial control limits by adding a center line based on the average of the five results and control limits based on the  $\overline{MR}$  from previous control charts for materials at the same nominal level. Set trial control limits for the *MR* chart nominal level.

8.7.2.4 After a minimum of 15 in-control data points are collected on the new material, perform an F test of sample variances for the new data set versus the historical variance demonstrated at nominal level of the new material. If the outcome of the F test is not significant then the precision estimate is updated by statistically pooling both sample variances. A significant F test should trigger an investigation for root cause(s).

8.7.2.5 Construct new *I* and *MR* charts (and optional *EWMA* overlay) for this new material as per Section 8, using the pooled  $\overline{MR}$ .

8.7.2.6 Switch over to the new I and MR charts upon depletion of current QC material.

8.7.3 Procedure 2-Q Procedure (see A1.9) (6):

8.7.3.1 This procedure is designed to alleviate the need for concurrent testing of two materials. A priori knowledge of the measurement process historical standard deviation applicable at the new QC material composition and property level is required.

NOTE 20—It is recommended that this standard deviation estimate be based on at least 50 data points.

8.7.3.2 When the Q procedure is operational (minimum of two data points), it can be used in conjunction with a MR chart constructed using the observations to provide QA of the measurement process.

NOTE 21—The Q procedure is not suitable for monitoring measurement system bias relative to an external value. It is designed to monitor the stability of the system mean. When used in conjunction with the MR chart, "in statistical control" status of the measurement system can be ascertained.

# 9. Evaluation of System Performance in Terms of Precision and Bias

9.1 Site Precision Estimated from Testing of QC Samples:9.1.1 Estimate the site precision of the measurement system at the level corresponding to a specific lot of QC sample as

2.46 times the  $\overline{MR}$  from the MR chart for that specific lot.

$$R' = 2.46 \times \overline{MR} \tag{3}$$

Note 22—The site precision standard deviation ( $\sigma_{R'}$ ) is estimated from the MR chart as  $R'/2.77 = \overline{MR}/1.128$ .

9.1.1.1 Alternatively, R' many be estimated using the rootmean-square formula for standard deviation:

$$\sigma_{R'} = \sqrt{\frac{\sum_{i=1}^{n} (I_i - \bar{I})^2}{n-1}}$$
(4)

$$R' = 2.77 \times \sigma_{R'} \tag{5}$$

9.1.1.2 For estimate of site precision standard deviation  $(\sigma_{R'})$  using retain results, first obtain the standard deviation of differences by applying the root-mean-square formula below to the differences between the original and retest results for samples with same nominal property level. If measurement process precision is known to be level independent, retest results from samples with different property levels can be used. Divide the standard deviation of differences by 1.414 to obtain the estimate for site precision standard deviation. ( $\sigma_{R'}$ ).

standard deviation of differences = (6)  

$$\sqrt{\frac{\sum (\text{individual difference - average difference})^2}{\text{total number of differences}}}$$

$$\sigma_{R'} = (\text{standard deviation of differences}) \div 1.414$$
 (7)

9.1.2 Compare R' to published reproducibility of the test method at the same level, if available. R' is expected to be less than or equal to the published value. Use the  $\chi^2$  test described in A1.7.

9.2 Measurement System Bias Estimated from Multiple Measurements of a Single Check Standard—If a minimum of 15 test results is obtained on a single check standard material under site precision conditions, then calculate the average of all the in-control individual differences plotted on the I chart. Perform a t test (see A1.6) to determine if the average is statistically different from zero.

9.2.1 If the outcome of the *t*test is that the average is not statistically different from zero, then the bias in the measurement process is negligible.

9.2.2 If the outcome of the t test is that the average is statistically different from zero, then the best estimate of the measurement process bias at the level of the check standard is the average.

9.3 Measurement System Bias Estimated from Measurements of Multiple Check Standards—When using multiple check standards, determine if there is a relationship between the bias and the measurement level.

9.3.1 Plot the pretreated results as per Section 8 versus their

corresponding ARVs. Examine the plot for patterns indicative of level-dependent bias.

9.3.2 If there is no discernible pattern, perform the t test as described in 9.2 to determine if the average of all the pretreated differences plotted on the I chart is statistically different from zero.

9.3.2.1 If the outcome of the t test is that the average is not statistically different from zero, then the bias in the measurement process is negligible.

9.3.2.2 If the outcome of the t test is that the average is statistically different from zero, then there is evidence that the measurement system is biased. The bias may be level dependent. However, the statistical methodology for estimating the bias/level relationship is beyond the scope of this practice.

9.3.3 If there is a discernible pattern in the plot in 9.3.1, then the measurement system may exhibit a level dependent bias. The statistical methodology for estimating the bias/level relationship is beyond the scope of this practice.

9.3.4 If a bias is detected in 9.3.2.2, or if the plot in 9.3.3 exhibits discernible patterns, investigate for root cause(s).

# 10. Validation of System Performance Estimates Using VA Samples

10.1 If the users decide to include VA sample testing as part of their QA program, then they should periodically evaluate the results obtained on the VA samples. The purpose of the evaluation is to establish whether the system performance estimates described in Section 9 are reasonably applicable to routinely tested samples.

10.2 VA sample test results should be evaluated independently through an internal or external audit system, or both. It is recommended that the internal audit team not be limited to the operators of the measurement system and their immediate supervisors.

10.3 Insofar as possible, analyze the results obtained on the VA samples separately and in the same manner as those from the routine QC and check standard testing program.

10.4 Using F or t tests, or both (see A1.8 and A1.6), statistically compare the system performance estimates obtained from the VA sample testing program to the measurement system accuracy and precision estimates from the QC sample testing program.

10.5 If the comparison reveals that the two estimates of the measurement system performance are not statistically equivalent, there is cause for concern that the actual performance of the measurement system may be significantly worse than estimated. Investigate thoroughly for the assignable cause(s) of this inconsistency, and eliminate it. Until the causes are identified and eliminated, the lab precision estimates of Section 9 should be considered suspect.

### ANNEX

#### (Mandatory Information)

#### A1. STATISTICAL QUALITY CONTROL TOOLS

#### A1.1 Purpose of this Annex

A1.1.1 The purpose of this annex is to provide guidance to practitioners, including worked examples, for the proper execution of the statistical procedures described in this practice.

#### A1.2 Pretreatment of Test Results (8.1 to 8.2.4)

A1.2.1 Throughout this annex,  $\{Y_i:i=1, ...,n\}$  denotes a sequence of as measured test results.  $\{I_i:i=1, ...,n\}$  will signify a sequence of test results after pretreatment, if necessary.

A1.2.2 If  $\{Y_i: i=1...n\}$  is a sequence of results from a single QC sample, then

$$I_i = Y_i \tag{A1.1}$$

with no pretreatment being required.

A1.2.2.1 An example of a sequence of results,  $Y_i$ , from a single QC sample is given in Columns 2 and 4 of Table A1.1.

A1.2.3 If  $\{Y_i:i=1...n\}$  is a sequence of results from a single check standard, from multiple check standards having nominally the same ARV, or from multiple check standards having different ARVs where the precision of the measurement system does not vary with level, and if  $\{X_i:i=1...n\}$  is the sequence of corresponding ARVs, then

$$I_i = Y_i - X_i \tag{A1.2}$$

The reproducibility standard deviation of the measurement process must be essentially the same for all values  $\{X_i\}$ .

A1.2.3.1 An example of a sequence of results from a single check standard is given in Table A1.2. The preprocessed result,  $I_i$ , is given in Column 4 of Table A1.2.

A1.2.4 If  $\{Y_i\}$  is a sequence of results from different check standards, and if the reproducibility varies with the level of the accepted reference values,  $\{X_i\}$ , then

$$I_i = (Y_i - X_i)/\sigma_i \tag{A1.3}$$

where  $\sigma_i$  are estimates of the reproducibility standard deviation of the measurement process at levels  $\{X_i\}$ .

TABLE A1.1 Example of a Sequence of Results from a Single QC Sample

Sequence Number <i>i</i>	QC/Check Standard Result $Y_i = I_i$	Sequence Number <i>i</i>	QC/Check Standard Result $Y_i = I_i$
1	55.3	14	55.2
2	55.8	15	56.5
3	56.3	16	55.7
4	56.1	17	55.6
5	55.8	18	55.2
6	55.5	19	55.7
7	55.3	20	56.1
8	55.4	21	56.3
9	56.6	22	55.2
10	56.1	23	55.4
11	55.0	24	55.4
12	55.5	25	55.6
13	55.5		

TABLE A1.2	Example of a Sequence of Results from a Single
	Check Standard

Sequence Number	Check Standard Result	Accepted Reference Value	Difference Result - ARV
	$(Y_i)$	$(ARV = X_i)$	$I_i$
1	55.3	55.88	-0.58
2	55.8	55.88	-0.08
3	56.3	55.88	0.42
4	56.1	55.88	0.22
5	55.8	55.88	-0.08
6	55.5	55.88	-0.38
7	55.3	55.88	-0.58
8	55.4	55.88	-0.48
9	56.6	55.88	0.72
10	56.1	55.88	0.22
11	55.0	55.88	-0.88
12	55.5	55.88	-0.38
13	55.5	55.88	-0.38
14	55.2	55.88	-0.68
15	56.5	55.88	0.62
16	55.7	55.88	-0.18
17	55.6	55.88	-0.28
18	55.2	55.88	-0.68
19	55.7	55.88	-0.18
20	56.1	55.88	0.22
21	56.3	55.88	0.42
22	55.2	55.88	-0.68
23	55.4	55.88	-0.48
24	55.4	55.88	-0.48
25	55.6	55.88	-0.28

A1.2.4.1 Table A1.3 shows an example of results for multiple check standards where the precision of the measurement system is level dependent.

TABLE A1.3 Example of Results for Multiple Check Standards Where the Precision of the Measurement System Is Level Dependent

Result Sequence Number, <i>i</i>	Raw Result Y <sub>i</sub>	$\operatorname{ARV}_{X_i}$	Raw Difference	$\sigma_i$	Preprocessed Result <i>I<sub>i</sub></i>
1	71.0	71.4	-0.40	1.14	-0.35
2	65.8	64.9	0.90	1.10	0.82
3	70.3	70.2	0.10	1.13	0.09
4	66.2	67.7	-1.50	1.11	-1.35
5	93.8	93.4	0.40	1.26	0.32
6	102.9	104.0	-1.10	1.33	-0.83
7	102.2	101.8	0.40	1.31	0.30
8	103.2	103.9	-0.70	1.32	-0.53
9	100	99.8	0.20	1.30	0.15
10	71.6	71.5	0.10	1.14	0.09
11	76.7	76.4	0.30	1.16	0.26
12	61.2	61.8	-0.60	1.08	-0.56
13	44.1	43.9	0.20	0.98	0.20
14	69.71	69.7	0.01	1.13	0.01
15	59.5	59.19	0.31	1.06	0.29
16	99.63	98.87	0.76	1.30	0.59
17	93.7	95.21	-1.51	1.27	-1.19
18	103.77	103.94	-0.17	1.32	-0.13
19	96.18	96.7	-0.52	1.28	-0.41
20	99.7	100.65	-0.95	1.31	-0.73
21	84.32	84.15	0.17	1.21	0.14
22	83.29	83.75	-0.46	1.21	-0.38
23	65.16	65.93	-0.77	1.10	-0.70
24	68.19	68.0	0.19	1.12	0.17

### A1.3 The Run Chart

A1.3.1 A run chart is a plot of results in chronological order that can be used to screen data for unusual patterns. Preferably, pretreated results are plotted. Use a run chart to screen data for unusual patterns such as continuous trending in either direction, unusual clustering, and cycles. Several non-random patterns are described in control chart literature. When control parameters have been added to a run chart, it becomes a control chart of individual values (*I* chart).

A1.3.2 Plot results on the chart. Plot the first result at the left, and plot each subsequent point one increment to the right of its predecessor. The points may be connected in sequence to facilitate interpretation of the run chart.

A1.3.3 Allow sufficient space in the *x*-axis direction to accommodate as many results as should be obtained from a consistent batch of material. Allow enough space in the *y*-axis direction to accommodate the expected minimum and maximum of the data.

A1.3.4 *Example of a Run Chart for QC Results*—The first 15 results from Column 2 of Table A1.1 are plotted in sequence as they are collected as shown in Fig. A1.1. The data would be examined for unusual patterns.

A1.3.5 *Example of a Run Chart for Multiple Results from a Single Check Standard*—The first 15 preprocessed results (differences) from Column 4 of Table A1.2 are plotted in sequence as they are collected as shown in Fig. A1.2. The data would be examined for unusual patterns.

A1.3.6 Example of a Run Chart for Results from Multiple Check Standards—The first 15 preprocessed results (differences scaled by  $\sigma_i$ ) from Table A1.3 are plotted in sequence as they are collected as shown in Fig. A1.3. The data would be examined for unusual patterns.

## A1.4 Normality Checks

A1.4.1 A normal probability plot (a special case of a q-q plot) is used to test the assumption that the observations are normally distributed. Since the control chart and limits prescribed in this practice are based on the assumption that the data behavior is adequately modeled by the normal distribution, it is recommended that a test of this normality assumption be conducted.

A1.4.1.1 To construct a normal probability plot:

(1) Create a column of the observations sorted in ascending order.

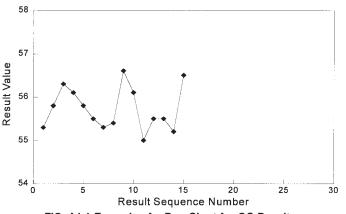


FIG. A1.1 Example of a Run Chart for QC Results

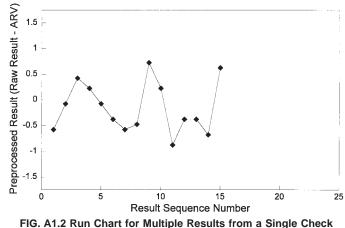


FIG. A1.2 Run Chart for Multiple Results from a Single Check Standard

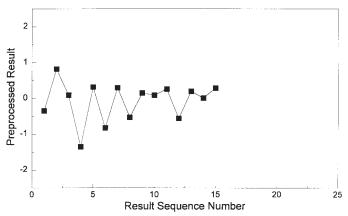


FIG. A1.3 Run Chart for Results from Multiple Check Standards

(2) Select the appropriate column from Fig. A1.4, based on the number of observations (n).

(3) Plot each observation in the sorted column (y-value) against its corresponding value from Fig. A1.4 (z-value).

A1.4.1.2 Visually inspect the plot for an approximately linear relationship. If the results are normally distributed, the plot should be approximately linear. Major deviations from linearity are an indication of nonnormal distributions of the differences.

NOTE A1.1—The assessment methodology of the normal probability plot advocated in this practice is strictly visual due to its simplicity. For statistically more rigorous assessment techniques, users are advised to consult a statistician.

A1.4.2 Anderson-Darling Statistic—The Anderson-Darling statistic is used to test for normality. The test involves the following steps:

A1.4.2.1 Order the non-outlying results such that  $x_1 \le x_2 \le \dots x_n$ 

A1.4.2.2 Obtain standardized variate from the  $x_i$  values as follows:

$$w_i = (x_i - \bar{x})/s \tag{A1.4}$$

for (i=1...n), where *s* is sample standard deviation of the results, and  $\bar{x}$  is the average of the results.

A1.4.2.3 Convert the  $w_i$  values to standard normal cumulative probabilities  $p_i$  values using the cumulative probability table for the standardized normal variate z (see Fig. A1.5):

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	n																	
Order	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30	31	32
1	-1.83	-1.86	-1.89	-1.91	-1.94	-1.96	-1.98	-2.00	-2.02	-2.04	-2.05	-2.07	-2.09	-2.10	-2.11	-2.13	-2.14	-2.15
2	-1.28	-1.32	-1.35	-1.38	-1.41	-1.44	-1.47	-1.49	-1.51	-1.53	-1.55	-1.57	-1.59	-1.61	-1.63	-1.64	-1.66	-1.68
3	-0.97	-1.01	-1.05	-1.09	-1.12	-1.15	-1.18	-1.21	-1.23	-1.26	-1.28	-1.30	-1.32	-1.35	-1.36	-1.38	-1.40	-1.42
4	-0.73	-0.78	-0.82	-0.86	-0.90	-0.93	-0.97	-1.00	-1.03	-1.05	-1.08	-1.10	-1.13	-1.15	-1.17	-1.19	1	
5	-0.52	-0.58	-0.63	-0.67	-0.72	-0.76	-0.79	-0.83	-0.86	-0.89	-0.92	-0.94	-0.97	-0.99	-1.01	-1.04		-1.08
6	-0.34	-0.40	-0.46	-0.51	-0.55	-0.60	-0.64	-0.67	-0.71	-0.74	-0.77	-0.80	-0.83	-0.85	-0.88	-0.90		
7	-0.17	-0.24	-0.30	-0.36	-0.41	-0.45	-0.50	-0.54	-0.58	-0.61	-0.64	-0.67	-0.70	-0.73	-0.76	-0.78		
8	0.00	-0.08	-0.15	-0.21	-0.27	-0.32	-0.37	-0.41	-0.45	-0.49	-0.52	-0.56	-0.59	-0.62	-0.65	-0.67	-0.70	
9	0.17	0.08	0.00	-0.07	-0.13	-0.19	-0.24	-0.29	-0.33	-0.37	-0.41	-0.45	-0.48	-0.51	-0.54	-0.57		-0.63
10	0.34	0.24	0.15	0.07	0.00	-0.06	-0.12	-0.17	-0.22	-0.26	-0.31	-0.34	-0.38	-0.41	-0.45		-0.51	
11	0.52	0.40	0.30	0.21	0.13	0.06	0.00	-0.06	-0.11	-0.16	-0.20		-0.28	-0.32	-0.35		-0.42	
12	0.73	0.58	0.46	0.36	0.27	0.19	0.12	0.06	0.00	-0.05	-0.10	-0.15	-0.19	-0.23	-0.26		-0.33	
13	0.97	0.78	0.63	0.51	0.41	0.32	0.24	0.17	0.11	0.05	0.00	-0.05	-0.09	-0.13	-0.17	-0.21	-0.25	
14	1.28	1.01	0.82	0.67	0.55	0.45	0.37	0.29	0.22	0.16	0.10	0.05	0.00	-0.04	-0.09	-0.13	0.10	
15	1.83	1.32	1.05	0.86	0.72	0.60	0.50	0.41	0.33	0.26	0.20	0.15	0.09	0.04	0.00	-0.04	0.00	
16		1.86	1.35	1.09	0.90	0.76	0.64	0.54	0.45	0.37	0.31	0.24	0.19	0.13	0.09	0.04	0.00	
17			1.89	1.38	1.12	0.93	0.79	0.67	0.58	0.49	0.41	0.34	0.28	0.23	0.17	0.13	0.08	0.04
18				1.91	1.41	1.15	0.97	0.83	0.71	0.61	0.52	0.45	0.38	0.32	0.26	0.21	0.16	0.12
19					1.94	1.44	1.18	1.00	0.86	0.74	0.64	0.56	0.48	0.41	0.35	0.30	0.25	0.20
20						1.96	1.47	1.21	1.03	0.89	0.77	0.67	0.59	0.51	0.45	0.39	0.33	0.28
21							1.98	1.49	1.23	1.05	0.92	0.80	0.70	0.62	0.54	0.48	0.42	0.36
22								2.00	1.51	1.26	1.08	0.94	0.83	0.73	0.65	0.57	0.51	0.45
23									2.02	1.53	1.28	1.10	0.97	0.85	0.76	0.67	0.60	0.53
24										2.04	1.55	1.30	1.13	0.99	0.88	0.78	0.70	0.63
25											2.05	1.57	1.32	1.15	1.01	0.90	0.81	0.72
26												2.07	2.09	1.35	1.17	1.04	0.93	0.83
27													2.09	2.10	1.36	1.19	1.06	0.95
28 29														2.10	2.11	1.58	1.21	1.08
30															2.11	2.13	1.40	1.23
31																2.13	2.14	
31																	2.14	2.15
32																		2.15

FIG. A1.4 z-Values

(A1.5)

$$p_i = \text{Probability} (z < w_i)$$

A1.4.2.4 Compute  $A^2$  as:

$$A^{2} = -\frac{\sum_{i=1}^{n} (2i-1) \left[ \ln(p_{i}) + \ln(1-p_{n+1-i}) \right]}{n} - n \qquad (A1.6)$$

A1.4.2.5 Compute  $A^{2*}$  as:

$$A^{2*} = A^{2} \left( 1 + \frac{0.75}{n} + \frac{2.25}{n^{2}} \right)$$
(A1.7)

A1.4.2.6 If the computed value of  $A^{2*}$  exceeds 0.752, then the hypothesis of normality is rejected for a 5 % level test.

A1.4.3 *Example of Normal Probability Plot for QC Results*—Once 15 results have been obtained (Table A1.1), they are sorted in ascending order and paired with the corresponding *z*-values from Fig. A1.4. The paired results (see Table A1.4) are plotted as (x, y) points (see Fig. A1.6). A line can be added to the plot to facilitate examination of the data for deviations from linearity.

A1.4.3.1 For the above example, the  $w_i$  and  $p_i$  values used in

the calculation of the Anderson-Darling statistic are shown in Table A1.4, as is the individual terms in the summation for  $A^2$ . The value for  $A^2$  is 0.415, and the value for  $A^{2*}$  is 0.440. Since this value is less than 0.752, the hypothesis of normality is accepted at the 95 % confidence level.

A1.4.4 Example of Normal Probability Plot for Multiple Results from a Single Check Standard—The first 15 preprocessed results (Table A1.2, Column 4) are sorted in ascending order and paired with the corresponding z-values from Fig. A1.4. The paired results (Table A1.5) are plotted as x, y points (Fig. A1.7). A line can be added to the plot to facilitate examination of the data for deviations from linearity.

A1.4.4.1 For this example, the  $w_i$ , and  $p_i$  values used in the calculation of the Anderson-Darling statistic are shown in Table A1.6, as are the individual terms in the summation for  $A^2$ . The value for  $A^2$  is 0.415, and the value for  $A^{2*}$  is 0.440. Since this value is less than 0.752, the hypothesis of normality is accepted at the 95 % confidence level.

A1.4.5 Example of Normal Probability Plot for Results from

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Order	33	34	35	36	37	38	39	40	41	42	43	44	45	46	47	48	49	50
1	-2.17	-2.18	-2.19	-2.20	-2.21	-2.22	-2.23	-2.24	-2.25	-2.26	-2.27	-2.28	-2.29	-2.29	-2.30	-2.31	-2.32	-2.33
	-1.69	-1.70	-1.72	-1.73	-1.74	-1.76	-1.77	-1.78	-1.79	-1.80	-1.81	-1.82	-1.83	-1.84	-1.85	-1.86	-1.87	-1.88
3	-1.43	-1.45	-1.47	-1.48	-1.49	-1.51	-1.52	-1.53	-1.55	-1.56	-1.57	-1.58	-1.59	-1.60	-1.61	-1.62	-1.64	-1.64
4	-1.25	-1.26	-1.28	-1.30	-1.31	-1.33	-1.34	-1.36	-1.37	-1.38	-1.40	-1.41	-1.42	-1.43	-1.44	-1.45	-1.47	-1.48
5	-1.10	-1.12	-1.13	-1.15	-1.17	-1.18	-1.20	-1.21	-1.23	-1.24	-1.26		-1.28	-1.29	-1.31	-1.32	-1.33	-1.34
6	-0.97	-0.99	-1.01	-1.02	-1.04	-1.06	-1.08	-1.09	-1.11	-1.12	-1.14	-1.15	-1.16	-1.18	-1.19	-1.20	-1.21	-1.23
7	-0.85	-0.87	-0.89	-0.91	-0.93	-0.95	-0.97	-0.98	-1.00	-1.02	-1.03	-1.05	-1.06	-1.07	-1.09	-1.10	-1.11	-1.13
8	-0.75	-0.77	-0.79	-0.81	-0.83	-0.85	-0.87	-0.89	-0.90	-0.92	-0.94	-0.95	-0.97	-0.98	-1.00	-1.01	-1.02	-1.04
9	-0.65	-0.67	-0.70	-0.72	-0.74	-0.76	-0.78	-0.80	-0.82	-0.83	-0.85	-0.87	-0.88	-0.90	-0.91	-0.93	-0.94	-0.95
10	-0.56	-0.58	-0.61	-0.63	-0.65	-0.67	-0.69	-0.71	-0.73	-0.75	-0.77	-0.79	-0.80	-0.82	-0.83	-0.85	-0.86	-0.88
11	-0.47	-0.50	-0.52	-0.55	-0.57	-0.59	-0.62	-0.64	-0.66	-0.67	-0.69	-0.71	-0.73	-0.74	-0.76	-0.78	-0.79	-0.81
12	-0.39	-0.42	-0.44	-0.47	-0.49	-0.52	-0.54	-0.56	-0.58	-0.60	-0.62	-0.64	-0.66	-0.67	-0.69	-0.71	-0.72	-0.74
13	-0.31	-0.34	-0.37	-0.39	-0.42	-0.44	-0.47	-0.49	-0.51	-0.53	-0.55	-0.57	-0.59	-0.61	-0.63	-0.64	-0.66	-0.67
14	-0.23	-0.26	-0.29	-0.32	-0.35	-0.37	-0.40	-0.42	-0.44	-0.46	-0.48	-0.50	-0.52	-0.54	-0.56	-0.58	-0.60	-0.61
15	-0.15	-0.19	-0.22	-0.25	-0.27	-0.30	-0.33	-0.35	-0.38	-0.40	-0.42	-0.44	-0.46	-0.48	-0.50	-0.52	-0.54	-0.55
16	-0.08	-0.11	-0.14	-0.17	-0.20	-0.23	-0.26	-0.29	-0.31	-0.33	-0.36	-0.38	-0.40	-0.42	-0.44	-0.46	-0.48	-0.50
17	0.00	-0.04	-0.07	-0.10	-0.14	-0.17	-0.19	-0.22	-0.25	-0.27	-0.30	-0.32	-0.34	-0.36	-0.38	-0.40	-0.42	-0.44
18	0.08	0.04	0.00	-0.03	-0.07	-0.10	-0.13	-0.16	-0.18	-0.21	-0.24	-0.26	-0.28	-0.30	-0.33	-0.35	-0.37	-0.39
19	0.15	0.11	0.07	0.03	0.00	-0.03	-0.06	-0.09	-0.12	-0.15	-0.18	-0.20	-0.22	-0.25	-0.27	-0.29	-0.31	-0.33
20	0.23	0.19	0.14	0.10	0.07	0.03	0.00	-0.03	-0.06	-0.09	-0.12	-0.14	-0.17	-0.19	-0.21	-0.24	-0.26	-0.28
21	0.31	0.26	0.22	0.17	0.14	0.10	0.06	0.03	0.00	-0.03	-0.06	-0.09	-0.11	-0.14	-0.16	-0.18	-0.21	-0.23
22	0.39	0.34	0.29	0.25	0.20	0.17	0.13	0.09	0.06	0.03	0.00	-0.03	-0.06	-0.08	-0.11	-0.13	-0.15	-0.18
23	0.47	0.42	0.37	0.32	0.27	0.23	0.19	0.16	0.12	0.09	0.06	0.03	0.00	-0.03	-0.05	-0.08	-0.10	-0.13
24	0.56	0.50	0.44	0.39	0.35	0.30	0.26	0.22	0.18	0.15	0.12	0.09	0.06	0.03	0.00	-0.03	-0.05	-0.08
25	0.65	0.58	0.52	0.47	0.42	0.37	0.33	0.29	0.25	0.21	0.18	0.14	0.11	0.08	0.05	0.03	0.00	-0.03
26	0.75	0.67	0.61	0.55	0.49	0.44	0.40	0.35	0.31	0.27	0.24	0.20	0.17	0.14	0.11	0.08	0.05	0.03
27	0.85	0.77	0.70	0.63	0.57	0.52	0.47	0.42	0.38	0.33	0.30	0.26	0.22	0.19	0.16	0.13	0.10	0.08
28	0.97	0.87	0.79	0.72	0.65	0.59	0.54	0.49	0.44	0.40	0.36	0.32	0.28	0.25	0.21	0.18	0.15	0.13
29	1.10	0.99	0.89	0.81	0.74	0.67	0.62	0.56	0.51	0.46	0.42	0.38	0.34	0.30	0.27	0.24	0.21	0.18
30	1.25	1.12	1.01	0.91	0.83	0.76	0.69	0.64	0.58	0.53	0.48	0.44	0.40	0.36	0.33	0.29	0.26	0.23
31	1.43	1.26	1.13	1.02	0.93	0.85	0.78	0.71	0.66	0.60	0.55	0.50	0.46	0.42	0.38	0.35	0.31	0.28
32	1.69	1.45	1.28	1.15	1.04	0.95	0.87	0.80	0.73	0.67	0.62	0.57	0.52	0.48	0.44	0.40	0.37	0.33
33	2.17	1.70	1.47	1.30	1.17	1.06	0.97	0.89	0.82	0.75	0.69	0.64	0.59	0.54	0.50	0.46	0.42	0.39
34		2.18	1.72	1.48	1.31	1.18	1.08	0.98	0.90	0.83	0.77	0.71	0.66	0.61	0.56	0.52	0.48	0.44
35			2.19	1.73	1.49	1.33	1.20	1.09	1.00	0.92	0.85	0.79	0.73	0.67	0.63	0.58	0.54	0.50
36				2.20	1.74	1.51	1.34	1.21	1.11	1.02	0.94	0.87	0.80	0.74	0.69	0.64	0.60	0.55
37					2.21	1.76	1.52	1.36	1.23	1.12	1.03	0.95	0.88	0.82	0.76	0.71	0.66	0.61
38						2.22	1.77	1.53	1.37	1.24	1.14	1.05	0.97	0.90	0.83	0.78	0.72	0.67
39							2.23	1.78	1.55	1.38	1.26	1.15	1.06	0.98	0.91	0.85	0.79	0.74
40								2.24	1.79	1.56	1.40	1.27	1.16	1.07	1.00	0.93	0.86	0.81
41									2.25	1.80	1.57	1.41	1.28	1.18	1.09	1.01	0.94	0.88
42										2.26	1.81	1.58	1.42	1.29	1.19	1.10	1.02	0.95
43											2.27	1.82	1.59	1.43	1.31	1.20	1.11	1.04
44												2.28	1.83	1.60	1.44	1.32	1.21	1.13
45													2.29	1.84	1.61	1.45	1.33	1.23
46														2.29	1.85	1.62	1.47	1.34
47															2.30	1.86	1.64	1.48
48																2.31	1.87	1.64
49																	2.32	1.88
50																		2.33
							FIG	A1.4 z-	Values	(		_						

FIG. A1.4 z-Values (continued)

*Multiple Check Standards*—The first 15 preprocessed results (Table A1.3, Column 6) are sorted in ascending order and paired with the corresponding *z*-values from Fig. A1.4. The paired results (Table A1.7) are plotted as *x*, *y* points (Fig. A1.8). A line can be added to the plot to facilitate examination of the

### data for deviations from linearity.

## A1.5 The Control Chart

A1.5.1 *I Chart*—The *I* chart is a run chart to which control limits and center line have been added. To establish placement

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	-0.09	-0.08	-0.07	-0.06	-0.05	-0.04	-0.03	-0.02	-0.01	0.00
-3.5	0.0002	0.0002	0.0002	0.0002	0.0002	0.0002	0.0002	0.0002	0.0002	0.0002
-3.4	0.0002	0.0003	0.0003	0.0003	0.0003	0.0003	0.0003	0.0003	0.0003	0.0003
-3.3	0.0003	0.0004	0.0004	0.0004	0.0004	0.0004	0.0004	0.0005	0.0005	0.0005
-3.2	0.0005	0.0005	0.0005	0.0006	0.0006	0.0006	0.0006	0.0006	0.0007	0.0007
-3.1	0.0007	0.0007	0.0008	0.0008	0.0008	0.0008	0.0009	0.0009	0.0009	0.0010
-3.0	0.0010	0.0010	0.0011	0.0011	0.0011	0.0012	0.0012	0.0013	0.0013	0.0013
-2.9	0.0014	0.0014	0.0015	0.0015	0.0016	0.0016	0.0017	0.0018	0.0018	0.0019
-2.8	0.0019	0.0020	0.0021	0.0021	0.0022	0.0023	0.0023	0.0024	0.0025	0.0026
-2.7	0.0026	0.0027	0.0028	0.0029	0.0030	0.0031	0.0032	0.0033	0.0034	0.0035
-2.6	0.0036	0.0037	0.0038	0.0039	0.0040	0.0041	0.0043	0.0044	0.0045	0.0047
-2.5	0.0048	0.0049	0.0051	0.0052	0.0054	0.0055	0.0057	0.0059	0.0060	0.0062
-2.4	0.0064	0.0066	0.0068	0.0069	0.0071	0.0073	0.0075	0.0078	0.0080	0.0082
-2.3	0.0084	0.0087	0.0089	0.0091	0.0094	0.0096	0.0099	0.0102	0.0104	0.0107
-2.2	0.0110	0.0113	0.0116	0.0119	0.0122	0.0125	0.0129	0.0132	0.0136	0.0139
-2.1	0.0143	0.0146	0.0150	0.0154	0.0158	0.0162	0.0166	0.0170	0.0174	0.0179
-2.0	0.0183	0.0188	0.0192	0.0197	0.0202	0.0207	0.0212	0.0217	0.0222	0.0228
-1.9	0.0233	0.0239	0.0244	0.0250	0.0256	0.0262	0.0268	0.0274	0.0281	0.0287
-1.8	0.0294	0.0301	0.0307	0.0314	0.0322	0.0329	0.0336	0.0344	0.0351	0.0359
-1.7	0.0367	0.0375	0.0384	0.0392	0.0401	0.0409	0.0418	0.0427	0.0436	0.0446
-1.6	0.0455	0.0465	0.0475	0.0485	0.0495	0.0505	0.0516	0.0526	0.0537	0.0548
-1.5	0.0559	0.0571	0.0582	0.0594	0.0606	0.0618	0.0630	0.0643	0.0655	0.0668
-1.4	0.0681	0.0694	0.0708	0.0721	0.0735	0.0749	0.0764	0.0778	0.0793	0.0808
-1.3	0.0823	0.0838	0.0853	0.0869	0.0885	0.0901	0.0918	0.0934	0.0951	0.0968
-1.2	0.0985	0.1003	0.1020	0.1038	0.1056	0.1075	0.1093	0.1112	0.1131	0.1151
-1.1	0.1170	0.1190	0.1210	0.1230	0.1251	0.1271	0.1292	0.1314	0.1335	0.1357
-1	0.1379	0.1401	0.1423	0.1446	0.1469	0.1492	0.1515	0.1539	0.1562	0.1587
-0.9	0.1611	0.1635	0.1660	0.1685	0.1711	0.1736	0.1762	0.1788	0.1814	0.1841
-0.8	0.1867	0.1894	0.1922	0.1949	0.1977	0.2005	0.2033	0.2061	0.2090	0.2119
-0.7	0.2148	0.2177	0.2206	0.2236	0.2266	0.2296	0.2327	0.2358	0.2389	0.2420
-0.6	0.2451	0.2483	0.2514	0.2546	0.2578	0.2611	0.2643	0.2676	0.2709	0.2743
-0.5	0.2776	0.2810	0.2843	0.2877	0.2912	0.2946	0.2981	0.3015	0.3050	0.3085
-0.4	0.3121	0.3156	0.3192	0.3228	0.3264	0.3300	0.3336	0.3372	0.3409	0.3446
-0.3	0.3483	0.3520	0.3557	0.3594	0.3632	0.3669	0.3707	0.3745	0.3783	0.3821
-0.2	0.3859	0.3897	0.3936	0.3974	0.4013	0.4052	0.4090	0.4129	0.4168	0.4207
-0.1	0.4247	0.4286	0.4325	0.4364	0.4404	0.4443	0.4483	0.4522	0.4562	0.4602
0.0	0.4641	0.4681	0.4721	0.4761	0.4801	0.4840	0.4880	0.4920	0.4960	0.5000

NOTE—Probability $(z < w_i)$ , where $w_i$	is the	sum of	f the	number	in the	left	column	and	top row.
F	IG. A1	.5 p <sub>i</sub> Va	alue	s					

positions of the control limits for the *I* chart, an estimate of the process variability will need to be obtained from the data. While there are several statistical techniques that can be used for this purpose, this practice advocates use of an *MR* of two chart for its simplicity and robustness to outliers. Produce an I chart only after a minimum of 15 preprocessed results have been obtained from the measurement system, and the data have been screened (see 8.3.1 and 8.3.2) and tested for normality (see A1.4).

A1.5.1.1 A horizontal center line is added at the level of the mean of all the results,  $\bar{I}$ :

$$\bar{I} = \frac{\sum_{i=1}^{n} I_i}{n} \tag{A1.8}$$

A1.5.1.2 Upper and lower control limits are added, also, computed from the  $\overline{MR}$  of two:

$$\overline{MR} = \frac{\sum_{i=1}^{n-1} |I_{i+1} - I_i|}{n-1}$$
(A1.9)

$$UCL_I = \bar{I} + 2.66 \,\overline{MR} \tag{A1.10}$$

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	0.00	0.01	0.02	0.03	0.04	0.05	0.06	0.07	0.08	0.09
0.0	0.5000	0.5040	0.5080	0.5120	0.5160	0.5199	0.5239	0.5279	0.5319	0.5359
0.1	0.5398	0.5438	0.5478	0.5517	0.5557	0.5596	0.5636	0.5675	0.5714	0.5753
0.2	0.5793	0.5832	0.5871	0.5910	0.5948	0.5987	0.6026	0.6064	0.6103	0.6141
0.3	0.6179	0.6217	0.6255	0.6293	0.6331	0.6368	0.6406	0.6443	0.6480	0.6517
0.4	0.6554	0.6591	0.6628	0.6664	0.6700	0.6736	0.6772	0.6808	0.6844	0.6879
0.5	0.6915	0.6950	0.6985	0.7019	0.7054	0.7088	0.7123	0.7157	0.7190	0.7224
0.6	0.7257	0.7291	0.7324	0.7357	0.7389	0.7422	0.7454	0.7486	0.7517	0.7549
0.7	0.7580	0.7611	0.7642	0.7673	0.7704	0.7734	0.7764	0.7794	0.7823	0.7852
0.8	0.7881	0.7910	0.7939	0.7967	0.7995	0.8023	0.8051	0.8078	0.8106	0.8133
0.9	0.8159	0.8186	0.8212	0.8238	0.8264	0.8289	0.8315	0.8340	0.8365	0.8389
1.0	0.8413	0.8438	0.8461	0.8485	0.8508	0.8531	0.8554	0.8577	0.8599	0.8621
1.1	0.8643	0.8665	0.8686	0.8708	0.8729	0.8749	0.8770	0.8790	0.8810	0.8830
1.2	0.8849	0.8869	0.8888	0.8907	0.8925	0.8944	0.8962	0.8980	0.8997	0.9015
1.3	0.9032	0.9049	0.9066	0.9082	0.9099	0.9115	0.9131	0.9147	0.9162	0.9177
1.4	0.9192	0.9207	0.9222	0.9236	0.9251	0.9265	0.9279	0.9292	0.9306	0.9319
1.5	0.9332	0.9345	0.9357	0.9370	0.9382	0.9394	0.9406	0.9418	0.9429	0.9441
1.6	0.9452	0.9463	0.9474	0.9484	0.9495	0.9505	0.9515	0.9525	0.9535	0.9545
1.7	0.9554	0.9564	0.9573	0.9582	0.9591	0.9599	0.9608	0.9616	0.9625	0.9633
1.8	0.9641	0.9649	0.9656	0.9664	0.9671	0.9678	0.9686	0.9693	0.9699	0.9706
1.9	0.9713	0.9719	0.9726	0.9732	0.9738	0.9744	0.9750	0.9756	0.9761	0.9767
2.0	0.9772	0.9778	0.9783	0.9788	0.9793	0.9798	0.9803	0.9808	0.9812	0.9817
2.1	0.9821	0.9826	0.9830	0.9834	0.9838	0.9842	0.9846	0.9850	0.9854	0.9857
2.2	0.9861	0.9864	0.9868	0.9871	0.9875	0.9878	0.9881	0.9884	0.9887	0.9890
2.3	0.9893	0.9896	0.9898	0.9901	0.9904	0.9906	0.9909	0.9911	0.9913	0.9916
2.4	0.9918	0.9920	0.9922	0.9925	0.9927	0.9929	0.9931	0.9932	0.9934	0.9936
2.5	0.9938	0.9940	0.9941	0.9943	0.9945	0.9946	0.9948	0.9949	0.9951	0.9952
2.6	0.9953	0.9955	0.9956	0.9957	0.9959	0.9960	0.9961	0.9962	0.9963	0.9964
2.7	0.9965	0.9966	0.9967	0.9968	0.9969	0.9970	0.9971	0.9972	0.9973	0.9974
2.8	0.9974	0.9975	0.9976	0.9977	0.9977	0.9978	0.9979	0.9979	0.9980	0.9981
2.9	0.9981	0.9982	0.9982	0.9983	0.9984	0.9984	0.9985	0.9985	0.9986	0.9986
3.0	0.9987	0.9987	0.9987	0.9988	0.9988	0.9989	0.9989	0.9989	0.9990	0.9990
3.1	0.9990	0.9991	0.9991	0.9991	0.9992	0.9992	0.9992	0.9992	0.9993	0.9993
3.2	0.9993	0.9993	0.9994	0.9994	0.9994	0.9994	0.9994	0.9995	0.9995	0.9995
3.3	0.9995	0.9995	0.9995	0.9996	0.9996	0.9996	0.9996	0.9996	0.9996	0.9997
3.4	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9998
3.5	0.9998	0.9998	0.9998	0.9998	0.9998	0.9998	0.9998	0.9998	0.9998	0.9998

FIG. A1.5 p<sub>i</sub> Values (continued)

$$LCL_I = \bar{I} - 2.66 \,\overline{MR} \tag{A1.11}$$

A1.5.1.3 Individual values that are outside the upper or lower control limits are indications of an unstable system, and efforts should be made to determine the cause. Optionally, any one of the following occurrences should be considered as potential signs of instability:

(1) Two out of three consecutive results on the I chart that are more than  $1.77 \overline{MR}$  distant from the center line in the same direction;

(2) Five consecutive results on the *I* chart that are more than 0.89  $\overline{MR}$  distant from the center line in the same direction;

(3) Eight or more consecutive points in the I chart that fall on the same side of the center line.

A1.5.2 *MR Chart*—A MR of two chart is obtained by plotting the sequential range of two values given by:

$$MR_i = |I_i - I_{i-1}| \tag{A1.12}$$

and connecting each point.

A1.5.2.1 The upper control limit for the MR chart is given by:

$$UCL_{MR} = 3.27 \ \overline{MR} \tag{A1.13}$$

A1.5.2.2 There is no lower control limit for an *MR* chart. A1.5.3 *EWMA Overlay*—A *EWMA* overlay is a trend line constructed from *EWMA* values calculated using the *I*-values. The *EWMA* trend line is typically overlaid on the *I* chart to enhance its sensitivity in detecting mean shifts that are small

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TABLE A1.4 Example Data for a Normal Probability Plot for QC Results

		Results	•		
Original Sequence No., <i>I</i>	z-value	Sorted Result	W <sub>i</sub>	p <sub>i</sub>	/ <sup>th</sup> Term in Eq A1.6
11	-1.83	55.0	-1.47	0.07	-5.91
14	-1.28	55.2	-1.07	0.14	-14.35
1	-0.97	55.3	-0.86	0.19	-18.70
7	-0.73	55.3	-0.86	0.19	-21.94
8	-0.52	55.4	-0.66	0.25	-25.77
6	-0.34	55.5	-0.46	0.32	-21.44
12	-0.17	55.5	-0.46	0.32	-25.34
13	0.00	55.5	-0.46	0.32	-22.80
2	0.17	55.8	0.15	0.56	-16.52
5	0.34	55.8	0.15	0.56	-18.46
10	0.52	56.1	0.76	0.78	-11.50
4	0.73	56.1	0.76	0.78	-10.80
3	0.97	56.3	1.16	0.88	-8.65
15	1.28	56.5	1.57	0.94	-5.79
9	1.83	56.6	1.77	0.96	-3.25

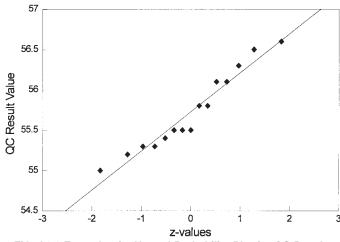


FIG. A1.6 Example of a Normal Probability Plot for QC Results

TABLE A1.5 Example Data for a Normal Probability Plot for Multiple Results from a Single Check Standard

Sort No.	Original Sequence No.	Sorted Result	z-value	W <sub>i</sub>	<i>P</i> <sub>i</sub>	<i>i</i> <sup>th</sup> Term in Eq A1.6
1	11	-0.88	-1.83	-1.47	0.07	-5.91
2	14	-0.68	-1.28	-1.07	0.14	-14.35
3	1	-0.58	-0.97	-0.86	0.19	-18.70
4	7	-0.58	-0.73	-0.86	0.19	-21.94
5	8	-0.48	-0.52	-0.66	0.25	-25.77
6	6	-0.38	-0.34	-0.46	0.32	-21.44
7	12	-0.38	-0.17	-0.46	0.32	-25.34
8	13	-0.38	0	-0.46	0.32	-22.80
9	2	-0.08	0.17	0.15	0.56	-16.52
10	5	-0.08	0.34	0.15	0.56	-18.46
11	10	0.22	0.52	0.76	0.78	-11.50
12	4	0.22	0.73	0.76	0.78	-10.80
13	3	0.42	0.97	1.16	0.88	-8.65
14	15	0.62	1.28	1.57	0.94	-5.79
15	9	0.72	1.83	1.77	0.96	-3.25

relative to the measurement system precision. Each *EWMA* value is a weighted average of the current result and previous results, with the weights decreasing exponentially with the age of the reading.

A1.5.3.1 A sequence of values, EWMA i, are calculated, and

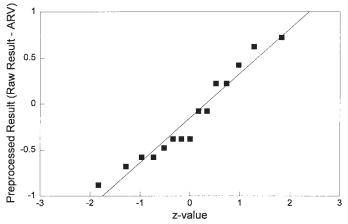


FIG. A1.7 Example of a Normal Probability Plot for Multiple Results from a Single Check Standard

TABLE A1.6 Example Data for a Normal Probability Plot for Results from Multiple Check Standards

			•			
Sort No.	Original Sequence No.	Sorted Result	<i>z</i> -value	W <sub>i</sub>	<i>P</i> <sub>i</sub>	<i>i</i> <sup>th</sup> Term in Eq A1.6
1	11	-1.83	-1.35	-3.12	0.00	-10.41
2	14	-1.28	-0.83	-1.94	0.03	-15.11
3	1	-0.97	-0.56	-1.32	0.09	-18.58
4	7	-0.73	-0.53	-1.25	0.11	-24.98
5	8	-0.52	-0.35	-0.84	0.20	-25.59
6	6	-0.34	0.01	-0.02	0.49	-19.68
7	12	-0.17	0.09	0.16	0.56	-19.93
8	13	0	0.09	0.16	0.56	-21.05
9	2	0.17	0.15	0.30	0.62	-22.33
10	5	0.34	0.2	0.41	0.66	-20.75
11	10	0.52	0.26	0.55	0.71	-11.91
12	4	0.73	0.29	0.62	0.73	-9.73
13	3	0.97	0.3	0.64	0.74	-9.99
14	15	1.28	0.32	0.69	0.75	-8.34
15	9	1.83	0.82	1.83	0.97	-1.02

overlaid on the *I* chart and connected. Use the following recursion equation:

$$EWMA_1 = I_1 \tag{A1.14}$$

$$EWMA_i = (1 - \lambda)EWMA_{i-1} + \lambda I_i$$
(A1.15)

where  $\lambda$  is the exponential weighting factor. For application of this practice, a  $\lambda$  value of 0.4 is recommended.

NOTE A1.2—For the *EWMA* trend, a  $\lambda$  value of 0.4 closely emulates the run rule effects of conventional control charts, while a value of 0.2 has optimal prediction properties for the next expected value. In addition, these  $\lambda$  values also conveniently places the control limits (3-sigma) for the *EWMA* trend at the 1 (for  $\lambda$ =0.2) to 1.5-sigma (for  $\lambda$ =0.4) values for *I* chart.

A1.5.3.2 The control limits for the *EWMA* chart are calculated using a weight ( $\lambda$ ) as follows:

$$UCL_{\lambda} = \overline{I} + 2.66 \ \overline{MR} \ \sqrt{\frac{\lambda}{2-\lambda}}$$
 (A1.16)

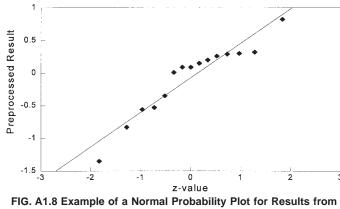
$$LCL_{\lambda} = \bar{I} - 2.66 \ \overline{MR} \sqrt{\frac{\lambda}{2 - \lambda}}$$
 (A1.17)

A1.5.4 Examples of Control Charts for QC and Check Standard Results:

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TABLE A1.7	Example Data for I Chart and EWMA Overlay for QC
	Results

	noe	Juito	
Sequence Number,	QC Result (Y <sub>i</sub> =I <sub>i</sub> )	Moving Range MR <sub>i</sub>	EWMA <sub>i</sub>
1	55.3		55.3
2	55.8	0.5	55.50
3	56.3	0.5	55.82
4	56.1	0.2	55.93
5	55.8	0.3	55.88
6	55.5	0.3	55.73
7	55.3	0.2	55.56
8	55.4	0.1	55.49
9	56.6	1.2	55.94
10	56.1	0.5	56.00
11	55	1.1	55.60
12	55.5	0.5	55.56
13	55.5	0.0	55.54
14	55.2	0.3	55.40
15	56.5	1.3	55.84
Average	55.73	0.500	
16	55.7	0.8	55.78
17	55.6	0.1	55.71
18	55.2	0.4	55.51
19	55.7	0.5	55.58
20	56.1	0.4	55.79
21	56.3	0.2	55.99
22	55.2	1.1	55.68
23	55.4	0.2	55.57
24	55.4	0.0	55.50
25	55.6	0.2	55.54

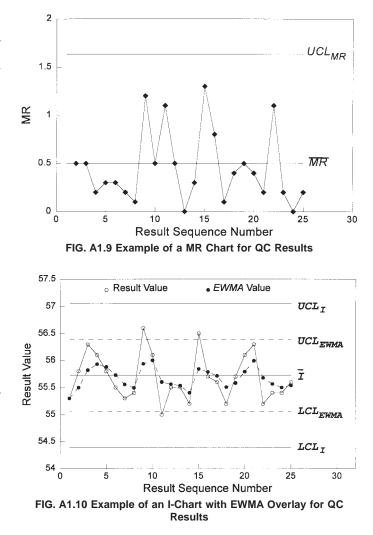


Multiple Check Standards

A1.5.4.1 Example of a MR Chart for QC Results—MR<sub>i</sub> values for the data from Table A1.1 are calculated and plotted in sequence. After 15 results are obtained, the  $\overline{MR} = 0.500$  value is calculated and added to the plot. Computations are shown in Table A1.7. A UCL<sub>MR</sub>=1.64 is added to produce the MR chart (Fig. A1.9).

A1.5.4.2 Example of I Chart and EWMA Overlay for QC Results—The average of the first 15 QC results (Table A1.7, Column 2) is calculated and plotted on the run chart as  $\overline{I}$  =55.73. The upper and lower control limits are calculated from Eq A1.10 and Eq A1.11 as 54.40 and 57.06 and added to the run chart to produce the I chart (Fig. A1.10). EWMA values (Table A1.7, Column 4) and EWMA control limits, 55.06 and 56.39, are overlaid on the I chart. Additional results and calculated EWMA values are added as they are determined.

A1.5.4.3 Example of a MR Chart for Multiple Results from a Single Check Standard— $MR_i$  values are calculated and



plotted in sequence. After 15 results are obtained (Table A1.2), the  $\overline{MR}$  value is calculated and added to the plot. A  $UCL_{MR}$  is added to produce the MR chart (see Fig. A1.11).

A1.5.4.4 Example of I Chart and EWMA Overlay for Multiple Results from a Single Check Standard—The average of the first 15 QC results (see Table A1.2, Column 4) is

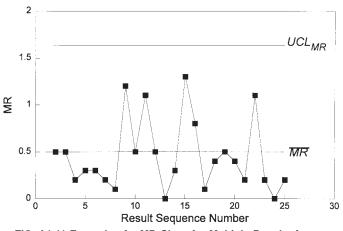


FIG. A1.11 Example of a MR Chart for Multiple Results from a Single Check Standard

calculated and plotted on the run chart as  $\overline{I}$ . The upper and lower control limits are calculated from Eq A1.6 and Eq A1.7 and added to the run chart to produce the *I* chart. *EWMA* values and *EWMA* control limits may be overlaid on the *I* chart (Fig. A1.12). Additional results and calculated *EWMA* values are added as they are determined. The *MR* values for this example are shown in Table A1.8, Column 3.)

A1.5.4.5 Example of a MR Chart for Results from Multiple Check Standards— $MR_i$  values are calculated and plotted in sequence. After 15 results are obtained (Table A1.3, Column 6, displayed again in Table A1.9), the  $\overline{MR}$  value is calculated and added to the plot. A  $UCL_{MR}$  is added to produce the MR chart (see Fig. A1.13).

A1.5.4.6 Example of I Chart and EWMA Overlay for Results from Multiple Check Standards—The average of the first 15 QC results (see Table A1.3, Column 6) is calculated and plotted on the run chart as  $\overline{I}$ . The upper and lower control limits are calculated from Eq A1.10 and Eq A1.11 and added to the run chart to produce the I chart. EWMA values and EWMA control limits may be overlaid on the I chart (Fig. A1.14). Additional results and calculated EWMA values are added as they are determined.

#### A1.6 t Test

A1.6.1 A two sided *t* test is used to check if a sample of values comes from a population with a mean different from an hypothesized value,  $\mu_0$ . In this practice, a *t* test may be performed on pretreated check standard test results to check for bias relative to the ARVs. Since during pretreatment, accepted reference value(s) have been subtracted from the raw results, the hypothesized mean value is zero.

A1.6.1.1 For the purpose of performing the t test, two methods for calculating the t value are presented:

(1) By the root-mean square method, the standard deviation of the pretreated results is calculated as:

$$S_I = \sqrt{\frac{\sum_{i=1}^{n} (I_i - \bar{I})^2}{n-1}}$$
 (A1.18)

(2)The *t* value is calculated as:

$$t = \sqrt{n}|\bar{I} - \mu_0| / S_I \tag{A1.19}$$

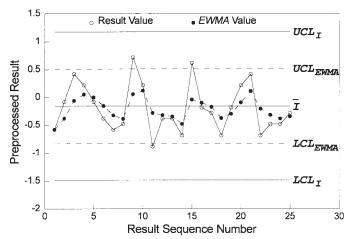


FIG. A1.12 Example of an I-Chart with EWMA Overlay for Multiple Results from a Single Check Standard

TABLE A1.8 Example Data for I Chart and EWMA Overlay for	
Multiple Results from a Single Check Standard	

· · ·			
Sequence Number	Check Standard Result ( <i>I</i> )	Moving Range, <i>MR</i> <sub>i</sub>	EWMA <sub>i</sub>
1	-0.58		-0.58
2	-0.08	0.5	-0.38
3	0.42	0.5	-0.06
4	0.22	0.2	0.05
5	-0.08	0.3	-0.00
6	-0.38	0.3	-0.15
7	-0.58	0.2	-0.32
8	-0.48	0.1	-0.39
9	0.72	1.2	0.06
10	0.22	0.5	0.12
11	-0.88	1.1	-0.28
12	-0.38	0.5	-0.32
13	-0.38	0.0	-0.34
14	-0.68	0.3	-0.48
15	0.62	1.3	-0.04
Average	-0.153	0.500	
16	-0.18	0.8	-0.10
17	-0.28	0.1	-0.17
18	-0.68	0.4	-0.37
19	-0.18	0.5	-0.30
20	0.22	0.4	-0.09
21	0.42	0.2	0.11
22	-0.68	1.1	-0.20
23	-0.48	0.2	-0.31
24	-0.48	0.0	-0.38
25	-0.28	0.2	-0.34

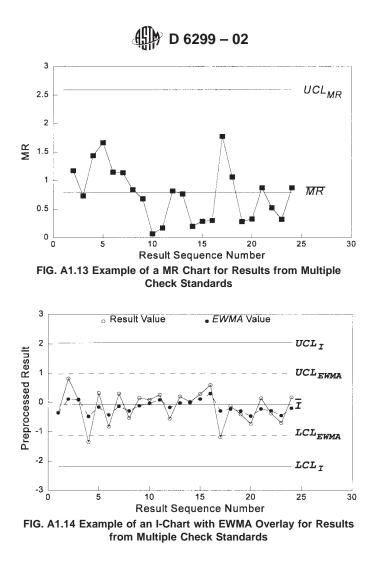
TABLE A1.9 Example Data for a MR Chart for Results from Multiple Check Standards

Result Sequence Number, <i>i</i>	Preprocessed Result, <i>I<sub>i</sub></i>	Moving Range, MR <sub>i</sub>	EWMA <sub>i</sub>
1	-0.35		-0.35
2	0.82	1.17	0.12
3	0.09	0.73	0.11
4	-1.35	1.44	-0.48
5	0.32	1.67	-0.16
6	-0.83	1.15	-0.43
7	0.30	1.13	-0.14
8	-0.53	0.83	-0.29
9	0.15	0.68	-0.12
10	0.09	0.06	-0.03
11	0.26	0.17	0.08
12	-0.56	0.82	-0.17
13	0.20	0.76	-0.02
14	0.01	0.19	-0.01
15	0.29	0.28	0.11
Average	-0.073	0.791	
16	0.59	0.3	0.30
17	-1.19	1.78	-0.29
18	-0.13	1.06	-0.23
19	-0.41	0.28	-0.30
20	-0.73	0.32	-0.47
21	0.14	0.87	-0.23
22	-0.38	0.52	-0.29
23	-0.7	0.32	-0.45
24	0.17	0.87	-0.20

where  $\mu_0$  is the hypothesized mean, which is zero (see A1.6.1).

(3) Alternatively, by the MR approach, compute the alternate t value as:

$$t_{MR} = \sqrt{n} |\bar{I} - \mu_0| / (\overline{MR}/1.128)$$
(A1.20)



where  $\mu_0$  is the hypothesized mean, which is zero (see A1.6.1).

freedom. If  $t_{MR}$  from Eq A1.20 is used, the appropriate degrees of freedom are (n-1)/2.

A1.6.1.2 Compare the computed t value from Eq A1.19 with the critical t values in Table A1.10 for (n-1) degrees of

	1 10	95th	Percentile	of	Student's	I fl	Distribution
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Degrees of Freedom	t
1	12.7062
2	4.3027
3	3.1824
4	2.7764
5	2.5706
6	2.4469
7	2.3646
8	2.3060
9	2.2622
10	2.2281
11	2.2010
12	2.1788
13	2.1604
14	2.1448
15	2.1314
16	2.1199
17	2.1098
18	2.1009
19	2.0930
20	2.0860
21	2.0796

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TABLE A1.10 Continued

TABLE A1.10 Continued					
Degrees of Freedom	t				
22	2.0739				
23	2.0687				
24	2.0639				
25	2.0595				
26	2.0555				
27	2.0503				
28	2.0484				
29	2.0452				
30	2.0423				
31	2.0395				
32	2.0369				
33	2.0345				
34	2.0322				
35	2.0301				
36	2.0281				
37	2.0262				
38	2.0244				
39	2.0244				
40	2.0211				
41	2.0195				
42	2.0181				
43	2.0167				
44	2.0154				
45	2.0141				
46	2.0129				
47	2.0117				
48	2.0106				
49	2.0096				
50	2.0086				
50					
55	2.0040				
60	2.0003				
65	1.9971				
70	1.9944				
75	1.9921				
80	1.99006				
85	1.98827				
90	1.98667				
95	1.98525				
100	1.98397				
105	1.98282				
110	1.98177				
	1.105.1				
115	1.98081				
120	1.97993				
125	1.97912				
130	1.97838				
135	1.97769				
140	1.97705				
145	1.97646				
150	1.97591				
155	1.97539				
160	1.97490				
165	1.97490				
170	1.97402				
175	1.97361				
180	1.97323				
185	1.97287				
190	1.97253				
195	1.97220				
200	1.97190				

A1.6.1.3 If the absolute value of the calculated t (or  $t_{MR}$ ) value is less than or equal to the critical t value, then  $\mu_0$  is statistically indistinguishable from the mean of the distribution. For the case of check standard testing, this would indicate that there is no statistically identifiable bias.

A1.6.1.4 If the absolute value of *t* is greater than the critical *t* value, then  $\mu_0$  is statistically distinguishable from the mean of the distribution, with 95 % confidence. For the case of check standard testing, this would indicate a statistically identifiable

bias in the measurement system.

A1.6.2 Example of t Test Applied to Multiple Results from a Single Check Standard—For the first 15 preprocessed results in Column 4 of Table A1.2,  $\bar{I}$  is -0.153. Since the results being analyzed are the difference relative to the ARV,  $\mu_0$  is zero. The standard deviation of the first 15 preprocessed results is 0.493, and the *t* value is 1.2034. The *t* value is less than the critical value for 14 degrees of freedom ( $t_{14}$ = 2.1448), so the average difference between the check standard results and the accepted

reference value is statistically indistinguishable from zero.

A1.6.3 *Example of t Test Applied to Results from Multiple Check Standards*—For the first 15 preprocessed results in Column 6 of Table A1.3,  $\overline{I}$  is -0.0719. Since the results being analyzed are the difference relative to the ARV,  $\mu_0$  is zero. The standard deviation of the first 15 preprocessed results is 0.550, and the *t* value is 0.506. The *t* value is less than the critical value for 14 degrees of freedom ( $t_{14} = 2.1448$ ), so the average difference between the check standard results and the accepted reference value is statistically indistinguishable from zero.

#### A1.7 Approximate Chi-Square Test

A1.7.1 The chi-square  $(\chi^2)$  test is used to compare the estimated site precision to a published reproducibility value, as instructed in 9.1.2.

A1.7.2 Compute the chi-square statistic:

$$\chi^2 = \frac{(n-1)R'^2}{2R^2}$$
(A1.21)

where R' is the estimated site precision ( $R'=2.46 \ \overline{MR}$ ) and R is the published reproducibility of the method.

A1.7.3 Compare the computed  $\chi^2$  value to the critical  $\chi^2$  value in Table A1.11, with (n-1)/2 degrees of freedom. If *n* is even, interpolate.

A1.7.3.1 If the computed  $\chi^2$  value exceeds the tabled value, then the site precision exceeds the published reproducibility of the method, with 95 % confidence.

A1.7.3.2 If the computed  $\chi^2$  value is less than or equal to the tabled value, then the site precision is either less than or statistically indistinguishable from the published reproducibility of the test method.

Degrees Freedom	X
7	14.1
8	15.5
9	16.9
10	18.3
11	19.7
12	21.0
13	22.4
14	23.7
15	25.0
16	26.3
17	27.6
18	28.9
19	30.1
20	31.4
21	32.7
22	33.9
23	35.2
24	36.4
25	37.7
26	38.9
27	40.1
28	41.3
30	43.8
35	49.8
40	55.8
45	61.7
50	67.5
60	79.1
70	90.5
80	101.9

A1.7.4 *Example*—The site precision calculated from R'= 2.46  $\overline{MR}$  for the first 15 QC results in Table A1.1 is 1.23. The published reproducibility for the measurement method at the 58.88 level is 1.05.  $\chi^2$  is therefore  $14 \cdot 1.23^2/2 \cdot 1.05^2 = 11.57$ . This value is less than the critical  $\chi^2$  value of 14.1 for 7 degrees of freedom, so the site precision is not statistically greater than the published reproducibility of the method.

#### A1.8 Approximate F Test

A1.8.1 In this practice, an approximate F test is used to compare the variation exhibited by a measurement system over two different time periods. It can also be used to compare the site precision estimated from a series of results from one QC sample with that estimated using a different QC sample (see 8.6.1).

A1.8.2 Compute the *F* value as:

$$F = \overline{MR}_{1}^{2} / \overline{MR}_{2}^{2}$$
(A1.22)

where  $\overline{MR}_1$  is the larger of the two average moving ranges, and  $\overline{MR}_2$  is the smaller.

A1.8.3 Compare the computed *F* value to the critical *F* value read from Table A1.12, with  $(n_1-1)/2$  degrees of freedom for the numerator and  $(n_2-1)/2$  degrees of freedom for the denominator.

A1.8.3.1 If the computed *F* value exceeds the tabled value, then the two precisions are statistically distinguishable. We can be 95 % confident that the process that gave rise to the moving range  $\overline{MR}_1$  is less precise (has larger site precision) than the process that produced  $\overline{MR}_2$ .

A1.8.3.2 If the computed F value is smaller than the tabled value, then the precisions of the two samplings of the measurement process are statistically indistinguishable.

NOTE A1.3—Although the approximate *F*-test is conducted at the 95 % probability level, the critical *F* values against which the calculated *F* is compared come from the 97.5 percentiles of the *F*-statistic. If the ratio  $\overline{MR}_a^2 / \overline{MR}_b^2$  is calculated without requiring that the larger variance is in the numerator, the ratio would have to be compared against both the lower 2.5 percentile point and the upper 97.5 percentile point of the *F*-distribution to determine if the two variances were statistically distinguishable. Because of the nature of the *F*-distribution, comparing  $\overline{MR}_a^2 / \overline{MR}_b^2$  to the 2.5 percentile point when  $\overline{MR}_a^2 / \overline{MR}_b^2$  is equivalent to comparing  $\overline{MR}_b^2 / \overline{MR}_a^2$  to the 97.5 percentile point. Requiring that larger variance is always in the numerator allows the" two-tailed" test to be accomplished in one step. If the variance of the two populations were equal, then there would be only a 2.5 % chance that  $\overline{MR}_1^2 > \overline{MR}_1^2$  by more than the tabulated amount, and a 2.5 % chance that  $\overline{MR}_2^2 > \overline{MR}_1^2$  by more than the tabulated amount with degrees of freedom reversed.

A1.8.4 If two precision estimates are statistically indistinguishable, they may be pooled into a single estimate. For example, if  $\overline{MR}_1$  was obtained from measurements on a single lot of QC sample material, while  $\overline{MR}_2$  was obtained from measurements on a different lot of material, and, if they are not statistically distinguishable, they may be pooled. The appropriate pooled precision estimate is

$$MR_{pooled} = \frac{(n_1 - 1)MR_1 + (n_2 - 1)MR_2}{n_1 + n_2 - 2}$$
(A1.23)

A1.8.5 *Example*—Table A1.13 contains QC results for a second QC sample measured by the same measurement system used to generate the results in Table A1.1. The  $\overline{MR}$  value for the 25 results from the original QC sample (Table A1.1) was 0.454.

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TABLE A1.12 97.5 Percentiles of the F Statistic

Denominator,	Numerator													
degrees of freedom	7	8	9	10	12	14	16	18	20	25	30	40	50	100
7	4.99	4.90	4.82	4.76	4.67	4.60	4.54	4.50	4.47	4.40	4.36	4.31	4.28	4.2
8	4.53	4.43	4.36	4.30	4.20	4.13	4.08	4.03	4.00	3.94	3.89	3.84	3.81	3.7
9	4.20	4.10	4.03	3.96	3.87	3.80	3.74	3.70	3.67	3.60	3.56	3.51	3.47	3.4
10	3.95	3.85	3.78	3.72	3.62	3.55	3.50	3.45	3.42	3.35	3.31	3.26	3.22	3.1
11	3.76	3.66	3.59	3.53	3.43	3.36	3.30	3.26	3.23	3.16	3.12	3.06	3.03	2.9
12	3.61	3.51	3.44	3.37	3.28	3.21	3.15	3.11	3.07	3.01	2.96	2.91	2.87	2.8
13	3.48	3.39	3.31	3.25	3.15	3.08	3.03	2.98	2.95	2.88	2.84	2.78	2.74	2.6
14	3.38	3.29	3.21	3.15	3.05	2.98	2.92	2.88	2.84	2.78	2.73	2.67	2.64	2.5
15	3.29	3.20	3.12	3.06	2.96	2.89	2.84	2.79	2.76	2.69	2.64	2.59	2.55	2.4
16	3.22	3.12	3.05	2.99	2.89	2.82	2.76	2.72	2.68	2.61	2.57	2.51	2.47	2.4
17	3.16	3.06	2.98	2.92	2.82	2.75	2.70	2.65	2.62	2.55	2.50	2.44	2.41	2.3
18	3.10	3.01	2.93	2.87	2.77	2.70	2.64	2.60	2.56	2.49	2.44	2.38	2.35	2.
19	3.05	2.96	2.88	2.82	2.72	2.65	2.59	2.55	2.51	2.44	2.39	2.33	2.30	2.
20	3.01	2.91	2.84	2.77	2.68	2.60	2.55	2.50	2.46	2.40	2.35	2.29	2.25	2.
25	2.85	2.75	2.68	2.61	2.51	2.44	2.38	2.34	2.30	2.23	2.18	2.12	2.08	2.
30	2.75	2.65	2.57	2.51	2.41	2.34	2.28	2.23	2.20	2.12	2.07	2.01	1.97	1.8
35	2.68	2.58	2.50	2.44	2.34	2.27	2.21	2.16	2.12	2.05	2.00	1.93	1.89	1.
40	2.62	2.53	2.45	2.39	2.29	2.21	2.15	2.11	2.07	1.99	1.94	1.88	1.83	1.
45	2.58	2.49	2.41	2.35	2.25	2.17	2.11	2.07	2.03	1.95	1.90	1.83	1.79	1.
50	2.55	2.46	2.38	2.32	2.22	2.14	2.08	2.03	1.99	1.92	1.87	1.80	1.75	1.0
60	2.51	2.41	2.33	2.27	2.17	2.09	2.03	1.98	1.94	1.87	1.82	1.74	1.70	1.0
70	2.47	2.38	2.30	2.24	2.14	2.06	2.00	1.95	1.91	1.83	1.78	1.71	1.66	1.
80	2.45	2.35	2.28	2.21	2.11	2.03	1.97	1.92	1.88	1.81	1.75	1.68	1.63	1.
90	2.43	2.34	2.26	2.19	2.09	2.02	1.95	1.91	1.86	1.79	1.73	1.66	1.61	1.
100	2.42	2.32	2.24	2.18	2.08	2.00	1.94	1.89	1.85	1.77	1.71	1.64	1.59	1.4

TABLE A1.13 Example of QC Results for a Second QC Sample Measured by the Same Measurement System

Number		MR	C <sub>n</sub>	LCL	UCL
1	54.2				
2	56.1	1.9	55.15	54.21	56.09
3	55.2	0.9	55.17	54.08	56.25
4	54.1	1.1	54.90	53.75	56.05
5	53.7	0.4	54.66	53.47	55.85
6	54	0.3	54.55	53.34	55.76
7	54.3	0.3	54.51	53.28	55.75
8	54.8	0.5	54.55	53.31	55.79
9	53.9	0.9	54.48	53.22	55.73
10	53.2	0.7	54.35	53.09	55.61
11	52.5	0.7	54.18	52.91	55.45
12	52.8	0.3	54.07	52.79	55.34
13	54.3	1.5	54.08	52.81	55.36
14	52.7	1.6	53.99	52.70	55.27
15	53.4	0.7	53.95	52.66	55.23
16	53.1	0.3	53.89	52.61	55.18
17	54	0.9	53.90	52.61	55.19
18	53.2	0.8	53.86	52.57	55.15
19	52.8	0.4	53.81	52.51	55.10
20	53.2	0.4	53.78	52.48	55.07
21	53.1	0.1	53.74	52.44	55.04
22	53.3	0.2	53.72	52.42	55.02
23	52.8	0.5	53.68	52.38	54.98

The  $\overline{MR}$  value for the 23 results for the new QC sample is 0.700. The *F* value is 2.38, which is less than the critical value of 3.33 for 11 and 12 degrees of freedom in the numerator and denominator, respectively. The precision of the measurements for the two QC batches is statistically indistinguishable.

### A1.9 Q-Procedure

A1.9.1 Collect and prepare a new batch of QC material while the current QC material supply remaining can support at least two additional analyses.

A1.9.2 Concurrently test the first sample of the new material with a routine analysis of the soon-to-be-depleted QC material.

Plot the result from the old material on its I chart, MR chart, EWMA chart, or Q chart, or a combination of these. If no special-cause signals are noted, then the result for the new material is considered to be valid.

A1.9.3 Plot the result from the new material as the first point on the Q chart.

NOTE A1.4—The Q chart is essentially a control chart of transformed statistics calculated from the conventional statistics normally plotted on control charts (for example, mean, range). This transformed statistics retains the information from the conventional statistics, but has the advantage of permitting plotting of all points on one standardized control chart.

A1.9.3.1 Center this value on the *y*-axis of the new chart. Scale the *y*-axis to allow room for the initial result plus and minus five historical standard deviations, where the standard deviations are appropriate to the level of the first result.

A1.9.3.2 No center line, nor upper or lower control limits, are plotted at this time.

A1.9.4 Subsequent QC sample testing may be done only on the new material.

A1.9.5 Plot subsequent QC results as points on the new Q chart. Do not connect the points.

A1.9.6 As each point (the  $n^{\text{th}}$  point) is plotted, compute and plot the center value and the upper and lower control limits applicable for this result.

A1.9.6.1 Center value,  $C_n = \sum_{i=1}^n I_i / n$ , where the sum includes the latest result,  $I_n$ . Optionally plot and connect the sequence of points  $\{C_n\}$  with a broken line. (Alternatively, replace any previous center line with a new line at the latest value of  $C_n$ .)

A1.9.6.2 Upper control limit,  $UCL_n = C_n + 3 \sigma \sqrt{(n-1)/n}$ , where  $\sigma$  is the historical standard deviation appropriate for test level  $C_n$ . For example, if the standard deviation is unchanged from the exhausted QC sample, then  $\sigma$ 

=  $\overline{MR}$  /1.128. Connect the sequence of points { $UCL_i$ } with a broken line. (Alternatively, replace any previous upper control limit lines with a new line at the latest  $UCL_n$ .

A1.9.6.3 Lower control limit,  $LCL_n = C_n - 3 \sigma \sqrt{(n-1)/n}$ . Connect the sequence of points  $\{LCL_i\}$  with a broken line. (Alternatively, replace any previous lower control limit lines with a new line at the latest  $LCL_n$ .)

A1.9.7 Individual values, current or earlier, which are outside the current upper or lower control limits, are indications of an unstable system, and efforts should be made to determine the cause. Optional run rules (corresponding to A1.5.1.3(1) to (3)) may also be applied to sequences of points using the current *UCL* and *LCL*, as early indicators of instability:

A1.9.7.1 Two consecutive results on the Q chart that are more than  $2\sigma\sqrt{(n-1)/n}$  distant from the current expected value,  $C_n$ , in the same direction;

A1.9.7.2 Five consecutive results on the *Q* chart that are more than  $\sigma \sqrt{(n-1)/n}$  distant from the current expected value in the same direction.

A1.9.7.3 Eight consecutive results on the Q chart that are on the same side of the current expected value.

A1.9.8 Continue or replace the MR chart, as appropriate.

A1.9.8.1 If the standard deviation for the new QC material is the same as for the old material, continue the old MR chart beginning with  $MR_2$ , that is, the second result from the new material.

A1.9.8.2 If the standard deviation appropriate to the level of the new material is different from the old, begin a new MR chart, starting with  $MR_2$ . The upper control limit for the new chart should be placed at  $3.69\sigma$ .

A1.9.8.3 After 15 results have been obtained with the new material, use a chi-square (see A1.7) or *F* test (see A1.8) to check that  $\sigma$  is appropriate for the new material.

A1.9.9 *EWMA* Overlay on a *Q* Chart—An *EWMA* chart may be overlaid on a *Q* chart, although it will not be meaningful until n > 5.

A1.9.9.1 The sequence of EWMA values,  $EWMA_i$ , are calculated, and overlaid on the *I* chart and connected. Use the following recursion:

$$EWMA_1 = I_1 \tag{A1.24}$$

$$EWMA_i = (1 - \lambda)EWMA_{i-1} + \lambda I_i$$
 (A1.25)

where  $\lambda$  is the exponential weighting factor, typically set to 0.4.

A1.9.9.2 The upper control limit for the EWMA chart is

$$UCL_{EWMA} = C_n + 3\sigma \sqrt{\left(\frac{\lambda}{2-\lambda}\right) + 2\left(\frac{1-\lambda}{2-\lambda}\right)(1-\lambda)^{2(n-1)} - \frac{1}{n}}$$
(A1.26)

A1.9.9.3 The lower control limit for the *EWMA* chart is

$$LCL_{EWMA} = C_n - 3\sigma \sqrt{\left(\frac{\lambda}{2-\lambda}\right) + 2\left(\frac{1-\lambda}{2-\lambda}\right)(1-\lambda)^{2(n-1)} - \frac{1}{n}}$$
(A1.27)

A1.9.10 *Example*—It is assumed that the collection of the QC results in Table A1.13 was started when there was sufficient quantity of QC batch 1 (Table A1.1) for two analyses. The individual values are plotted as they are collected, (squares in Fig. A1.15), and the  $C_n$  and  $UCL_n$  and  $LCL_n$  values are calculated and added for each new result. Recall that  $\overline{MR}$  from the first 15 measurements on batch 1 was 0.500. The new control limits (Table A1.13, Columns 5 and 6) are compared to the current and previous results. Note that, for this example, the second result is considered "out of control" when UCL is calculated. The "out-of-control" character of this result is confirmed as UCL is updated with additional data. The Q chart clearly shows that the results for the new QC sample trend downward with time.

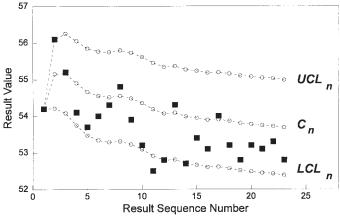


FIG. A1.15 Example of a Q-Chart for a New QC Sample



#### REFERENCES

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